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# An Experimental Investigation of Price Dispersion and Cycles* 

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#### Abstract

We report a continuous time experiment studying the classic Burdett and Judd (1983) model, whose unique Nash equilibrium (NE) features dispersed prices. Adaptive dynamics predict that the NE is stable for one parameter set we use, and unstable for another parameter set. The empirical price distribution turns out to be close to the NE distribution for the stable parameter set overall, but for the unstable parameter set the empirical distribution skews towards higher prices in its NE support interval. We offer an empirical definition of price cycles in terms of changes over time in robust measures of central tendency (median) and dispersion (interquartile range). By that definition, the data exhibit persistent cycles in both treatments, but with larger cycles for the unstable parameters. Results are roughly similar for professional and student sellers and for limited information treatments.


Keywords: price dispersion, laboratory experiment, cycles, stability of equilibrium
JEL Classifications: C73, C92, D83

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## 1 Introduction

Few laws are disobeyed more frequently than the Law of One Price (LOP). For a multitude of goods ranging from alphabet blocks to zenon software, at a given point in time one can find a broad range of prices for apparently identical units (e.g., Baye et al., 2004). Jensen (2007) illustrates the major consequences such price dispersion can have for market efficiency and welfare. At least since Stigler (1961), economists have pointed to search frictions as the underlying cause. The now-canonical equilibrium search model of Burdett and Judd (1983) obtains a dispersed price distribution as a (mixed strategy) best response by sellers to the search behavior of buyers, who in turn are best responding to the dispersed prices. Varian's (1980) search model likewise explains LOP violations in terms of a mixed strategy Nash equilibrium, albeit with less strategic buyers.

It is not clear how to take such equilibrium predictions to the data. Subtle heterogeneities (e.g., in shipping, reputation, or individualized discounts) may be unobservable and are difficult to account for properly. More fundamentally, field behavior typically consists of asynchronous price adjustments in real time, and the proper connection to an equilibrium mixture of prices is debatable.

Since Smith (1962), economists have used data from simple laboratory markets to help resolve ambiguities in interpreting field data. Such markets can directly implement the primitives of a theoretical model, and the model's prediction successes and failures can help guide subsequent analysis of data from more complex ongoing markets in the field. Here we use new tools to study price dispersion dynamics in experimental markets, varying the search frictions in a controlled manner while eliminating confounds such as subtle heterogeneities and synchronized price changes. Thus our setting is ideally suited for initial clean tests of dynamics and equilibrium in dispersed prices.

Broadly speaking, in a series of experimental markets inspired by the Burdett-Judd model, we find that the equilibrium distribution is not a bad predictor for the observed overall (time-collapsed) price distribution. However, we uncover a major role for dynamics. Individual sellers continually move their prices around within the dispersion. Beyond that,
the dispersion itself moves over time, and exhibits strong and persistent cycles that move consistently in the direction predicted by adaptive dynamics.

Our project also makes several methodological contributions of interest in their own right. We introduce nonparametric techniques for identifying cycles in price distributions; these may be more useful than standard spectral (or time series) techniques when cycle frequencies and amplitudes are irregular in lab or field data. Second, in contrast to the usual discrete period lab software which enforces simultaneous pricing decisions, we use continuous time software that enables asynchronous pricing decisions in real time. The importance of that distinction is exemplified in Wang (2009), who finds that introducing enforced simultaneous price setting for gasoline in Australia significantly changes behavior relative to the natural asynchronous setting.

Third, our treatments are multidimensional. In terms of search technology, we examine stable (S) versus unstable (U) search parameters. As explained below, S theoretically promotes convergence to equilibrium if sellers follow simple adaptive learning rules, while U theoretically prevents convergence. We also examine immediate versus speed-limited price changes, and high versus limited information on counterfactual profits and on other sellers' prices. A final treatment variation is in the subject pool. Some sessions use the usual pool of undergraduate college students, while others use professionals - executive MBAs, many of whom have experience in setting prices for their company's products. Those sessions constitute an artefactual field experiment (Harrison and List, 2004) and provide an anchor point for a bridge to future investigations of price dispersion and cycles using field data.

Before positioning our work within existing literature, we offer a few remarks on external validity. Our primary goal is to provide clean initial evidence on foundational issues regarding price dispersion and dynamics, giving theory its best shot; in the terminology of List (2020), ours is a WAVE1 study and establishes that the range of applicability is not empty. No single study, whether field or lab, can in itself guarantee generalization to any significantly different field or lab setting. However, we would argue that price competition gives strong basic incentives that will be the same in any other market which has the same general features, in particular, consumers who vary in their price sensitivity. Further, a study reporting robust
results can increase confidence that the range of applicability is broad. Thus our robust findings for multidimensional treatments - information conditions that bracket what might be seen in the wider world, as well as various adjustment speeds and experience levels of subjects - shows that the basic incentive structure continues to determine behavior and bodes well for generalizability.

The literature on price cycles goes back at least to Edgeworth (1925), who describes gradually declining prices in an undercutting phase and upward jumps after capacity constraints begin to bind. He regards the succession of phases as predictable, but not the time spent in each phase. Capacity constraints are not relevant to our work, and Edgeworth's cycles have little dispersion at most points of time, but the intuition may nevertheless be helpful.

Maskin and Tirole (1988) obtain equilibrium cycles in an alternating-moves Bertrand duopoly model on a discrete price grid. Their dynamic pricing model has many Markovperfect Nash equilibria, one of which is turn-taking as sellers undercut each other before returning to a high price. The resulting cycles are very regular with minimal dispersion. It is not clear how to generalize this equilibrium beyond two sellers. The equilibrium model of Deneckere and Kovenock (1992) includes capacity constraints, and price cycles are also featured in the general equilibrium literature on the Scarf example, e.g., Anderson et al. (2004), and Goeree and Lindsay (2016). None of these papers has a role for dispersed prices.

Considerable evidence exists for cyclical pricing behavior in real world oligopolistic markets; see, for example, Slade (1998), Noel (2007), Wang (2009) or Byrne and de Roos (2019). Some of this evidence is reminiscent of Edgeworth cycles, with episodes of slowly falling prices often followed by a large upward jumps.

Regarding equilibrium price dispersion, a crucial question is whether in practice the equilibrium is approachable. Given the strategic nature of the Burdett-Judd setting and the complexity of the equilibrium in question, this might seem unlikely. Nonetheless, Hopkins and Seymour (2002) show that simple adaptive learning rules can result in equilibrium play emerging over time, even in this setting. In fact, depending on exogenous parameters describing search efficacy, the Burdett-Judd equilibrium dispersion can be either stable or
unstable under adaptive learning. In the stable case, we would therefore predict convergence (at least at the population level) towards the equilibrium price distribution, but it is unclear what to expect in the unstable case.

Our experiment is guided by that prediction and open question, along with previous experimental work by ourselves and others. Cason and Friedman (2003) study the comparative static predictions of Burdett-Judd model with the best tools then available, that is, a discrete time experiment with enforced simultaneous moves. Further, only unstable parameter sets were used. Cason, Friedman and Wagener (2005) re-examine the same data with an eye on dynamics, and find suggestions of Edgeworth-like price cycles; e.g., small price changes are more likely to be downward than upward. Using this older, discrete time and simultaneous move experimental technology, however, only one or two complete cycles could be observed in a typical run of 30 trading periods. Morgan et al. (2006) take the Varian (1980) model to the lab and find support for its static predictions of price dispersion. Brown Kruse et al. (1994) earlier examined mixed strategy equilibrium in an oligopoly setting with quantity constraints similar to Edgeworth's.

To prepare for the present study, the current authors examined static equilibrium and dynamics in simple bimatrix games. Cason et al. (2010) uses standard discrete time protocols and finds that a point prediction called TASP outpredicts Nash equilibrium in a set of $4 \times 4$ bimatrix games. Cason et al. (2014) studies asynchronous real time choices and introduces new techniques for detecting cycles in symmetric $3 \times 3$ bimatrix (Rock-Paper-Scissors style) games. In the current study, we extend those techniques for cycle detection to much larger action spaces of a far more realistic and complex environment. It thus addresses questions about price dynamics and price dispersion that have intrinsic interest, as well as testing predictions of learning theory.

New technology for conducting continuous time experiments has gained favor recently, in part because it can enable long-run behavior to emerge in much less clock time than with standard discrete time periods, and in part because it provides sharper and richer data. Benndorf et al. (2016), Oprea et al. (2011) and Stephenson (2019) find substantial support for the long run stability predictions of evolutionary game theory, albeit for simple games
with only two alternative actions. Experiments with larger action sets are reported in Calford and Oprea (2017) and Stephenson and Brown (2017). However, we believe that ours is the first continuous time experiment of oligopolistic price competition, and the first to study cycles in dispersed prices. Importantly, the continuous time technology allows observation of hundreds of cycles per experimental session, instead of the small handful of cycles seen in discrete time.

The next section summarizes existing theory on equilibrium in the Burdett-Judd model and on its dynamic stability. A new discussion follows about the nature of cycles in high (or infinite) dimensional spaces, and the section concludes with a list of testable hypotheses. Section 3 details the experimental design, which features a novel videogame-like screen for subjects who serve as sellers in a six-firm oligopoly played in continuous time. Experimental treatments are explained, including stable (S) versus unstable (U) search parameter vectors, and price adjustment speeds (almost instantaneous, medium and slow) that potentially affect the nature of dynamic adjustment and convergence. The section also describes the subpool of professional subjects and the high versus limited information treatments.

Section 4 presents the results. At the aggregate level, we compare overall empirical distributions by treatment to the equilibrium distributions, and find impressively close agreement in some treatments (especially S parameters, student subjects and full information) and looser agreement in others. Regressions confirm several significant treatment effects on medians (our preferred measure of central tendency) and some marginally significant effects on dispersion (measured as the interquartile range, IQR). The section then documents the presence and persistence of cycles in (median, IQR) space, and reports support for most (but not all) of the testable hypotheses.

Section 5 summarizes and interprets the main findings, and suggests possible future research. Appendix A supplements theoretical analysis, Appendix B supplements the data analysis, and Appendix C presents the experiment instructions provided to subjects.

## 2 Theoretical Considerations

This section reviews equilibrium in the sequential search price dispersion model of Burdett and Judd (1983), hereafter BJ83, and the stability analysis in Hopkins and Seymour (2002), hereafter HS02. A discussion follows on how one might detect cycles in distributions supported on an interval of prices. The section concludes with a list of predictions to be tested.

### 2.1 Equilibrium

BJ83 explains how price dispersion - a range of prices for a homogeneous good - can be supported in equilibrium. A continuum of firms chooses prices and a continuum of consumers searches sequentially for low prices. An equilibrium is a distribution of prices such that all prices in its support earn the same (maximal) profit and such that consumer search behavior is optimal given the distribution.

Consumer search is both sequential and noisy in the BJ83 model. Each consumer initially receives a random number of price quotations, and then either buys at the cheapest amongst existing quotations or else sinks search cost $c>0$ to obtain a fresh sample of price quotations. The optimal policy is to adopt a particular reservation price $p^{*}$ and to search again if and only if no current price quote is equal to or below $p^{*}$.

Sellers thus face buyers who vary ex post in the quality of their information. Some buyers will randomly receive multiple quotations and will buy from a lower priced seller, while others will only see one price and will take it even if it is only slightly below $p^{*}$. This has two consequences. First, there is a non-trivial range of prices $[0, \underline{p})$ that sellers will never choose because these prices are dominated by $p^{*}$, a price at which demand is positive due to the positive fraction of buyers who see only one price. Second, demand is decreasing in price, because low prices attract relatively informed consumers as well as the lucky uninformed while high prices attract only unlucky, relatively uninformed consumers. With decreasing demand (and, for analytical convenience, zero production cost), sellers get the same payoff at a high price in $\left[\underline{p}, p^{*}\right]$ that yields a few highly profitable sales as they do
at a lower price that yields a larger volume of less profitable sales.

This leads to the main BJ83 result: given a positive fraction $q_{1}$ of consumers who see only one price, there exists a unique equilibrium. The relevant equal payoff condition yields a specific distribution of prices with continuous cumulative distribution function $F(p)$ with support interval $\left[\underline{p}, p^{*}\right]$. The main exogenous variables, which determine the shape and support of the equilibrium distribution, are the cost of search $c>0$ and the number of quotations received by consumers. In particular, assume that each consumer with probability $q_{k}$ receives $k=1,2$ or 3 iid quotations from $F$. Then the information conditions can be summarized by a point in the 2-D simplex, i.e., by a nonnegative vector $q=\left(q_{1}, q_{2}, q_{3}\right)$ with $q_{3}=1-q_{1}-q_{2}$.

It is not hard to see that a seller of rank $u \in[0,1]$ in the price distribution (the highest priced seller has rank $u=1$ ) faces demand

$$
\begin{equation*}
D(u)=\mu\left(q_{1}+2 q_{2}(1-u)+3 q_{3}(1-u)^{2}\right), \tag{1}
\end{equation*}
$$

where $\mu$ is the number of buyers per seller. Imposing the equal payoff condition (see Appendix A), one can then show that the inverse equilibrium distribution is

$$
\begin{equation*}
F^{-1}(u)=\frac{p^{*} q_{1}}{q_{1}+2 q_{2}(1-u)+3 q_{3}(1-u)^{2}}, \tag{2}
\end{equation*}
$$

on the support interval $\left[\underline{p}, p^{*}\right]$ with

$$
\begin{equation*}
\underline{p}=F^{-1}(0)=\frac{p^{*} q_{1}}{q_{1}+2 q_{2}+3 q_{3}} \text { and } p^{*}=F^{-1}(1)=\frac{c}{1-q_{1}} . \tag{3}
\end{equation*}
$$

Note that in equilibrium sellers are indifferent over choosing any price in the equilibrium support $\left[\underline{p}, p^{*}\right]$ since they all generate the same profit, $p^{*} \mu q_{1}$. Note also that in this model, all buyers consume a single indivisible unit of the good, so total demand is the same for all prices (below $p^{*}$ ); there is no welfare loss from demand restriction as in textbook models of monopoly. Higher prices simply imply higher transfers from buyers to sellers.

### 2.2 Stability

The model of BJ83 implies strategic interaction amongst sellers, with the best response of each seller depending on the distribution of prices amongst other sellers. In such complex situations of strategic uncertainty, it is unclear whether players realistically can play equilibrium in a one-shot setting. Evolutionary game theory and adaptive learning theory investigate whether equilibrium can be reached if players continually play the same game, adjusting their strategies over time using simple rules.

HS02 identify parameters for which this equilibrium distribution is dynamically stable or unstable under such evolutionary or adaptive dynamics. We can explain intuitively in terms of the relative size of wins, draws and losses. In our oligopoly games, a "win" is where a seller has the lowest price and attracts lots of customers $(D(0)$ in equation (1)), a "loss" is where a seller has the highest price and only can sell to buyers who receive a single price quote $(D(1))$, and a "draw" is where the seller has the median price amongst sellers $(D(0.5))$.

Mixed Nash equilibria are unstable in games for which draws are almost as good as wins because then learning dynamics tend to spend more time near pure states where prices are less dispersed and spend less time near the mixed equilibrium. By contrast, a mixed equilibrium tends to be stable where draw payoffs are only slightly better than loss payoffs and are far below win payoffs; here adaptive learners take on more dispersion in their prices to try to get wins. As shown in Appendix A, this analogy allows us to characterize stability in terms of the convexity of the demand function $D(u)$ given in equation (1). Simply put, convexity leads to very high sales from low prices and thus big win payoffs and so stability, whereas concavity implies a big loss penalty due to a significant reduction in sales from having the highest price and hence instability. Following HS02, the following result is stated in terms of PDA (positive definite adaptive) learning dynamics, a class that contains various forms of reinforcement learning and imitation and evolutionary dynamics as special cases.

Proposition 1. A dispersed price equilibrium is stable under all PDA learning dynamics if and only if $D(u)$ is log convex; it is unstable under all PDA learning dynamics if and only if $D(u)$ is log concave.


Figure 1: Regions of search parameter space for which the unique equilibrium of BJ83 is stable or unstable under PDA dynamics. The red (resp. blue) dot shows the parameter values $U$ (resp. $S$ ) used in the experiment. Adapted from HS02, Figure 1.

Figure 1 illustrates this result. Recall that $q=\left(q_{1}, q_{2}, q_{3}\right)$, a point in the 3 -simplex, summarizes the information available to consumers. The function $D(u)$ is $\log$ convex in the region labelled Stable near the $q_{1}=1$ corner, where losses and draws are less painful. In this region, all PDA dynamics converge to the unique Nash dispersed price equilibrium. In contrast, $D(u)$ is log concave in the region labeled Unstable, which includes a large neighborhood of the $q_{2}=1$ corner; here all PDA dynamics move away from the equilibrium. In the unlabeled intermediate region, $D(u)$ is neither $\log$ concave nor log convex and consequently the equilibrium could be stable under some PDA dynamics but unstable under others.

Our experiment employs two particular parameter vectors highlighted in the figure, and summarized in Table 1. These were chosen for their simplicity and focality: the $U$ vector was the principal case analyzed in Cason and Friedman (2003), and we chose the $S$ vector to be as similar as possible while lying well within the Stable region identified in Figure 1.

Unstable Parameter Vector: Red dot $U$ in Figure 1 at $q_{U}=(2 / 3,1 / 3,0)$; consumer search cost is $c=0.6$. Figure 2 shows the equilibrium price cumulative distribution function $F_{U}(p)=2-9 / 5 p$ on support interval $\left[\underline{p}, p^{*}\right]=[0.9,1.8]$. Here the mean price is 1.25 , and expected profits are 120 per seller.

Stable Parameter Vector: Blue dot $S$ in Figure 1 at $q_{S}=(0.8,0.1,0.1)$, with $c=0.36$. The equilibrium cdf is $F_{S}(p)=1-\left(\left(-p+\sqrt{p} \sqrt{24\left(p^{*}-p\right)+p}\right) / 3 p\right)$ on $\left[\underline{p}, p^{*}\right]=[1.11,1.8]$, with mean price 1.47 and expected profits 144 per seller.



Figure 2: Nash equilibrium price density functions (panel a) and cumulative distribution functions (panel b) for Stable (orange dashed) and Unstable (blue) parameters. A uniform distribution (green) is provided for contrast.

Figure 2 shows that the two equilibrium distributions are somewhat different, despite (by construction) having the same upper bound $p^{*}=1.8$. The unique mode of $F_{U}$ is at its lower bound $\underline{p}=0.9$, while $F_{S}$ has a mode at its upper bound as well as at its lower bound $\underline{p}=1.1$. Further, $F_{S}$ stochastically dominates $F_{U}$. Finally, the two equilibrium distributions are quite distinct from a simple uniform distribution which is also plotted in Figure 2 for contrast. Thus, equilibrium demands a very specific distribution of prices in each case.

### 2.3 Cycles

As explained in Appendix A, price choices of a continuum of agents at a given moment $t$ (or, in our experiment, mixed choices of a finite number of agents) are represented by a point $F^{t}$ in the space $\mathcal{F}$ of all cumulative distributions over the relevant price interval. Over time


Figure 3: Detecting limit cycles in high dimensions. Given a point $P$ on a stable limit cycle with tangent vector $N$, take the intersection of the hyperplane orthogonal to $N$ with a ball of radius $\varepsilon$ around $P$. This Poincaré section, $\mathrm{PS}=\{z \in B(P, \varepsilon): z \cdot N=0\}$, is the cycle detector.
$t>0$, agents adjust prices and the distribution $F^{t}$ moves around in that infinite dimensional space. The movement could be quite complicated, even chaotic. In such a high dimensional space, how could we tell if the price distribution eventually approached the equilibrium distribution, or if it converged to a nice limit cycle?

Figure 3 illustrates a classic technique, due to Poincaré (1890), for detecting cycles and convergence in finite ( $n \geq 2$ ) or infinite dimensional spaces. Suppose that we have a point $P$ that may be on a limit cycle, and a hyperplane through $P$ that is transversal to the cycle. (E.g., the hyperplane could be the vectors orthogonal to the tangent vector $N$ at $P$.) The intersection PS of that hyperplane with a ball of radius $\varepsilon$ around $P$ is called a Poincaré section, and it is our cycle detector.

The detector works as follows. At a sequence of times $0 \leq t_{1}<t_{2}<\ldots$, the evolving price distribution $\left\{F^{t}: t \geq 0\right\}$ intersects PS. At each such time $t_{k}$, assign the sign +1 if the price path goes in the predicted direction (e.g., if the tangent to $\left\{F^{t}: t \geq 0\right\}$ at $t=t_{k}$ has a positive dot product with $N$ ) and -1 if in the opposite direction. If, for $k=1,2, \ldots$, the
points $x_{k}=F^{t_{k}} \in \mathrm{PS}$ all have the same sign and converge to some point $F^{*} \in \mathrm{PS}$, then we have a (stable) limit cycle.

Appendix A explains how to deal with some technical issues in applying this method, but practical difficulties remain. How do we find a point $P$ on (or near) a limit cycle if there is noise in the system? Which dynamics should we use to compute the tangent vector at $P$ ? How big should $\varepsilon$ be? With noise, we would see too few observations $x_{k}$ if $\varepsilon$ is too small, but if it is too large we might pick up other pieces of a limit cycle, going the wrong way.

In the background is a metaphysical question: what is a cycle? Physicists and chemists traditionally think of cycles as oscillators with regular frequencies that can be analyzed nicely using spectral techniques. Oscillation may persist at a finite amplitude, or may be damped and ultimately disappear. By contrast, economists (and some biologists) tend to think of cycles as processes that never settle down, but that have distinct phases that recur in a particular sequence. They would likely exclude damped oscillations, but might include somewhat irregular recurrence as in some noisy or chaotic systems. ${ }^{1}$ We suspect that most economists would not recognize as cycles any process with subtle periodic movements in the price distribution that change neither the central location nor the degree of dispersion.

From that perspective, it makes practical as well as intuitive sense for economists to work, not with the infinite dimensional space $\mathcal{F}$, but rather with a two dimensional projection that tracks the central tendency of the price distribution and its degree of dispersion. Our data analysis therefore will look for cycles in median price (a robust measure of central tendency) and interquartile range ( $\mathrm{IQR}=75$ th minus 25 th percentile prices, a robust measure of dispersion). The obvious alternatives, mean and variance, are less robust to extreme behavior by an individual subject. Nevertheless, as we document in Appendix B, all of our qualitative conclusions regarding cycles still remain when we use mean and variance.

Figure 4 uses simulations to illustrate what we might see if players gradually adapt their prices towards the current best response. The blue line tracks the median price and

[^1]

Figure 4: Simulated (perturbed) BR dynamics. Sample trajectories under slightly smoothed best response dynamics for $U$ (blue) and $S$ (orange) parameters generate clockwise cycles in median price and interquartile range (IQR) space. Labeled dots show Nash equilibrium values.
the interquartile range as the distribution evolves under perturbed best response dynamics for $U$ parameters. (Specifically, the simulations use equation (16) in Appendix A: logitperturbed best response dynamics with precision parameter $\lambda=4$.) Starting at an initial distribution with median price near 1.45 and IQR near 0.55 , the dynamics for $U$ soon bring the median price and IQR down to near their NE values of (1.20, 0.41 ). But they don't stay there. Instead, the distribution's median and IQR cycle clockwise around a tilted trapezoid surrounding the NE. Starting at the top right corner of the trapezoid, prices fall as the best response is to undercut slightly, and the IQR decreases as prices cluster more tightly. Then, near the bottom left corner of the trapezoid, the best response jumps to 1.8 , the maximum price. Dispersion increases as more low-price sellers move quickly towards high prices. When most sellers charge high prices, dispersion again decreases, The blue line heads back towards the top right corner of the trapezoid, and then undercutting resumes.

The orange line in the same Figure shows, starting from a nearby initial distribution, a simulated time path for $S$ parameters, using the same perturbed best response dynamics. This time path indeed converges to $(1.48,0.37)$, the NE median and IQR.

Note that Edgeworth's best response cycle would not look the same. In discrete time, if all sellers were simply to play a best response to prices of the previous period, then prices would steadily descend in small steps from 1.8 to 0.9 in the $U$ case, or to 1.11 in the $S$ case, and then jump to 1.8 . Price dispersion would almost always be nearly zero.

### 2.4 Testable Hypotheses

Since a continuum of sellers cannot fit in a laboratory, we follow Cason and Friedman (2003) in implementing the Burdett and Judd model with a finite number of sellers, in this case, six. One can write $F(p)=u=1-\left(\frac{r_{i}-1}{5}\right)$ or $1-u=\left(\frac{r_{i}-1}{5}\right)$ where $r_{i} \in\{1,2, \ldots, 6\}$ is the rank of firm $i$ with the convention that the highest priced seller has rank 1 and the lowest has rank 6. Recall that demand $s\left(p_{i}\right)=0$ whenever $p_{i}>p^{*}=1.8$; to avoid trivialities we allow only choices $p_{i} \leq p^{*}$. Thus by equation (1), seller $i$ faces demand

$$
\begin{equation*}
d\left(p_{i}\right)=\mu\left(q_{1}+2 q_{2}\left(\frac{r_{i}-1}{5}\right)+3 q_{3} \max \left[0,\left(\frac{r_{i}-1}{5}\right)\left(\frac{r_{i}-2}{4}\right)\right]\right) \tag{4}
\end{equation*}
$$

where $\mu=100$ is the ratio of (automated) buyers to sellers in our experiment. If two or more sellers post exactly the same price, then demand is shared by those sellers. For example, if two sellers tie for lowest price then they both obtain the average of the sales of rank 6 and rank 5 . Since production costs are zero and $q_{U}=(2 / 3,1 / 3,0)$, sellers' payoff function in the $U$ treatment is

$$
\begin{equation*}
\pi_{U}\left(p_{i}, p_{-i}\right)=100 p_{i}\left(\frac{2}{3}+\frac{2}{3}\left(\frac{r_{i}-1}{5}\right)\right) \tag{5}
\end{equation*}
$$

and is

$$
\begin{equation*}
\pi_{S}\left(p_{i}, p_{-i}\right)=100 p_{i}\left(\frac{4}{5}+\frac{1}{5}\left(\frac{r_{i}-1}{5}\right)+\frac{3}{10} \max \left[0,\left(\frac{r_{i}-1}{5}\right)\left(\frac{r_{i}-2}{4}\right)\right]\right) \tag{6}
\end{equation*}
$$

in the $S$ treatment, where $q_{S}=(0.8,0.1,0.1)$. This 6 -player game approximates the BJ83 model; indeed, it is easy to see that its unique symmetric NE is for each player to use the mixed strategy described by $F_{U}(p)$ in the $U$ treatment and by $F_{S}(p)$ in the $S$ treatment. $^{2}$

[^2]Table 1: Nash Equilibrium (NE) predictions for Stable (S) and Unstable (U) parameters

| Variable | Symbol | $U$ | $S$ |
| :--- | :---: | :---: | :---: |
| Distribution of Price Quotes | $q$ | $(2 / 3,1 / 3,0)$ | $(0.8,0.1,0.1)$ |
| Search Cost | $c$ | 0.60 | 0.36 |
| Buyers per Seller | $\mu$ | 100 | 100 |
| Max NE Price | $p^{*}$ | 1.80 | 1.80 |
| Min NE Price | $p$ | 0.90 | 1.11 |
| Average NE Price | $m$ | 1.25 | 1.47 |
| Median NE Price | $m e d$ | 1.20 | 1.48 |
| NE Price Interquartile Range | $I Q R$ | 0.41 | 0.37 |
| NE Price Standard Deviation | $\sigma$ | 0.25 | 0.21 |
| Average Seller NE Profit | $\pi$ | 120 | 144 |

It has been suggested to us that the static NE is implausible in this setting, since the calculations are simply too demanding. We have some sympathy with this view, but point out that BJ83 is far from unique in relying on complex equilibria. To give the original theory a fair shot, our high information treatment gives subjects extensive information about prices and profits. All our treatments give subjects substantial opportunities to learn from experience, as explained in the next section.

If they do learn, then learning theory makes several predictions. First, as set out in Section 2.2, behavior in the stable and unstable settings should be very different. Second, our simulations of adaptive dynamics (Figure 4) show clockwise motion in (median, IQR) price space. Thus, we predict cycles in our experiment also to be clockwise and not counterclockwise. Finally, adaptive dynamics in theory can work in both high and low information settings, so there is no predicted effect from changing the information given to subjects.

Another consideration is that human subjects might try to collude, e.g., all sellers might simply charge $p^{*}$. This concern led to our decision to have groups of six, where collusion is more difficult than in smaller groups.

Initially we had another behavioral concern - in continuous time, subjects might make so many large and frequent price changes that the payoff landscape would seem a random blur, discouraging systematic adaptive learning or evolutionary dynamics. Therefore we include adjustment speed as treatment variable, as detailed in the next section. Along with
fast (virtually instantaneous) price change, we consider slow and medium speed limits, which may affect the focus and quality of subjects' learning processes. We have no sound theoretical basis for predicting how that treatment will affect overall price distributions or the strength of cycles, but the treatment clearly affects the minimum time to complete a cycle, and thus the maximum frequency.

These theoretical considerations lead to the following testable hypotheses.
I. Nash Equilibrium: For all subject pools (professionals and students), all information conditions and all adjustment speeds, the empirical distributions of prices will be stationary and equal to the NE distribution, $F_{S}(p)$ for Stable parameters and $F_{U}(p)$ for Unstable parameters. Table 1 gives numerical values.
II. Price cycles: For all subject pools, all information conditions and all adjustment speeds,
(a) Cycles will be clockwise in (median, IQR) space.
(b) Cycle amplitude will converge to 0 for $S$ parameters.
(c) Cycle amplitude will converge to a positive number bounded away from 0 for $U$ parameters.
(d) Cycle frequency will be highest for the Fast adjustment speed, and lowest for the Slow adjustment speed.

## 3 Experimental Design

In our within-subjects design, each session of the experiment has some market periods with the $S$ parameters and other periods with the $U$ parameters specified in Table 1. For example, $q_{U}=(2 / 3,1 / 3,0)$ means that drawing one price has twice the likelihood of drawing two prices with $U$ parameters, while with $S$ parameters $q_{S}=(0.8,0.1,0.1)$, a single price draw is very likely but drawing three prices is also possible. Given those parameters, in each period a continuum of buyers is simulated to behave optimally given the search costs and their price
draws. The continuum eliminates sampling variance, so current payoff is a deterministic functions of sellers' price profile.

Between 12 and 24 human subjects participate as sellers in each session. Each subject is assigned permanently to a market with six sellers, and each session has at least two such markets. Sellers are physically interspersed and spread widely throughout a 30 station lab with partitions that obscure subjects' identities and assignments.

We used only business professionals as sellers in 9 of our 39 markets. These subjects were 25 to 54 years old (average age 36), and averaged 13.4 years of work experience. Nearly half ( $42.4 \%$ ) report price-setting experience in their occupation. ${ }^{3}$ The remaining 30 markets ( 180 subjects) use only subjects from the usual pool of undergraduate students.

As illustrated in Figure 5 and in the instructions included in Appendix C, sellers select and adjust their prices in continuous time using a horizontal slider restricted to the interval [ $0,1.80]$. The corresponding profit flow given in equations $(5-6)$ is shown as the vertical height of the green circle above the current slider position.

In the "high information" condition, we provide two additional visual cues. Each seller sees differently-colored circles on the horizontal axis indicating the price choices (but not the profit flow) of the other five sellers in the market. This corresponds to the perfect price monitoring available in some markets (e.g., Byrne and de Roos, 2019). The same display also includes a black line showing the seller's own counterfactual profit flow at every feasible price, given the prices currently offered by the other sellers. The highest point on this line shows the player the best response to the current price distribution; the sawtooth shape seen in Figure 5 is typical. Of course, the line is constantly in motion as sellers continuously adjust their prices. The purpose is to give subjects an impression of the profit opportunities available to them, and we believe that our subjects find it far more useful than detailed mathematical formulas. Twenty-three of the 39 markets, including all 9 markets with professionals, provide this counterfactual profit (or "payoff landscape") line.

[^3]

Figure 5: Example Seller Price Choice and Feedback Screen

To study how such contemporary information may affect pricing behavior, we also conducted sessions where it is suppressed. In 8 "limited info-prices only" markets the display eliminates the payoff landscape line, but still shows other sellers' prices on the horizontal axis. Another 8 "limited info-no prices" markets eliminates both cues and provide no information about other sellers' current or past prices. An example screen is included in Appendix C. ${ }^{4}$

The two panels on the right of sellers' screens have elapsed time as the horizontal axis. The large lower panel shows the seller's accumulating profit as the green area swept out by the vertical height of the green circle. The narrow upper panel's line graphs show all six sellers' prices chosen so far; this panel is suppressed in the "no prices" condition.

Besides the primary S vs U parameter treatment, we have a secondary within-subject treatment, adjustment speed, with three alternative conditions. When a subject moves the price slider or clicks on a new target price, her actual price moves from its old value to the new target at a constant speed. In the Fast condition, the move is virtually instantaneous (0.072 seconds for a typical adjustment spanning $10 \%$ of the available price range). In the

[^4]Slow adjustment condition the same adjustment takes 3.6 seconds to complete, and in the Medium condition it takes 0.72 seconds.

Each session with undergraduate student subjects is divided into 32 trading periods. Initial prices are randomly assigned each period. Trading periods each last 90 seconds, except for Slow adjustment periods after the first three sessions. Those periods last 150 seconds, in order to permit more cycles.

The 32 trading periods in each student pool market are divided into eight blocks of four periods each. The within-subjects treatments (i.e., S vs. U parameters and price adjustment speed) are held constant within each four-period block and varied in a balance fashion across blocks in each session, as detailed in Table 18 in Appendix B. Due to time constraints for the business professionals, their sessions last only 16 periods so each market provides data for only 4 of the 6 within-subject treatment combinations. Since assignment of subjects to separate six-seller markets is permanent in each session, the 233 subjects provide 36 independent sets of observations for each treatment combination.

Sessions were conducted at Purdue University between October 2017 and February 2020. The student subject sessions took about 80 to 90 minutes and earnings per subject (including a $\$ 5.00$ participation payment) averaged $\$ 28.95$, with a standard deviation of $\$ 1.80$. We increased the exchange rate for the business professionals, so their 40 to 50 minute sessions' earnings averaged $\$ 44.24$ with a standard deviation of $\$ 2.31$.

## 4 Results

Our data consist of prices posted by subjects sampled at 400 ms intervals, i.e., 2.5 times per second. Initial prices in the first six markets were drawn randomly without replacement from the set $\{0.50,0.75, \ldots, 1.75\}$; the data analysis drops the first 15 seconds of each period when sellers are adjusting prices towards the narrower range where they typically operate for the rest of the period. In the remaining 33 markets we used the narrower set $\{1.10$, $1.20, \ldots, 1.60\}$ of initial prices, enabling us to reduce the excluded initial adjustment time to
the first 10 seconds of each period. The nonparametric tests reported below all employ the conservative convention that the unit of observation is the independent 6 -seller market.

### 4.1 Overview

The top panels of Figure 6 compare overall observed cumulative price distributions in Stable treatments to the Nash equilibrium of the BJ83 model. The prediction seems right on target except for a clump of prices right at the upper endpoint $p^{*}=1.80$ (about $9 \%$ of choices with high information but up to $14 \%$ with limited information), and a few prices a little below the predicted lower endpoint $\underline{p}=1.11$ of the predicted support interval. The corresponding Unstable NE prediction shown in the lower half of Figure 6 is almost first-order stochastically dominated by each of the empirical cumulative price distributions, due mainly to fewer low prices than predicted. In each panel, adjustment speed seems to have little impact.


Figure 6: Cumulative distributions of observed vs NE prices in Stable (top panels) and Unstable (bottom panels) treatments. Samples exclude first 10-15 seconds of each period.

Table 2 presents summary statistics for pooled subject pools, with other treatments
disaggregated. ${ }^{5}$ The upper left side shows that the median prices observed for the Stable treatment are remarkably close to the equilibrium prediction of 1.48 , and they are not significantly different from this benchmark in the high information condition. ${ }^{6}$ As predicted, median prices are lower in the Unstable treatment, but they nevertheless are significantly greater than the equilibrium prediction of 1.20 . The lower half of Table 2 indicates that median prices are greater when sellers have more limited price information and payoff feedback. This is consistent with the price CDFs shown on the right side of Figure 6, which have a larger mass of prices on the maximum price.

Table 2: Mean of Summary Price Statistics (and standard errors across periods)

| High Information (Pooled) | Median Price |  | Interquartile Range |  |  |  |  |  |  |
| :--- | :---: | :---: | :---: | ---: | :---: | :---: | :---: | :---: | :---: |
| Nash Equilibrium (Stable) | 1.48 |  | 0.37 |  |  |  |  |  |  |
| Stable-Fast Adjust | 1.482 | $(0.007)$ | $0.433^{* *}$ | $(0.006)$ |  |  |  |  |  |
| Stable-Medium Adjust | 1.496 | $(0.009)$ | $0.447^{* *}$ | $(0.009)$ |  |  |  |  |  |
| Stable-Slow Adjust | 1.469 | $(0.009)$ | $0.398^{*}$ | $(0.007)$ |  |  |  |  |  |
| Nash Equilibrium (Unstable) | 1.20 |  | 0.41 |  |  |  |  |  |  |
| Unstable-Fast Adjust | $1.324^{* *}$ | $(0.008)$ | 0.411 | $(0.009)$ |  |  |  |  |  |
| Unstable-Medium Adj. | $1.339^{* *}$ | $(0.011)$ | 0.403 | $(0.008)$ |  |  |  |  |  |
| Unstable-Slow Adjust | $1.317^{* *}$ | $(0.010)$ | 0.377 | $(0.008)$ |  |  |  |  |  |
| Limited Information (Pooled) | Median Price |  |  |  |  |  |  | Interquartile Range |  |
| Nash Equilibrium (Stable) | 1.48 |  | 0.37 |  |  |  |  |  |  |
| Stable-Fast Adjust | $1.528^{* *}$ | $(0.006)$ | $0.426^{* *}$ | $(0.007)$ |  |  |  |  |  |
| Stable-Medium Adjust | $1.550^{* *}$ | $(0.007)$ | $0.411^{* *}$ | $(0.006)$ |  |  |  |  |  |
| Stable-Slow Adjust | $1.548^{* *}$ | $(0.006)$ | 0.393 | $(0.006)$ |  |  |  |  |  |
| Nash Equilibrium (Unstable) | 1.20 |  |  | 0.41 |  |  |  |  |  |
| Unstable-Fast Adjust | $1.356^{* *}$ | $(0.009)$ | $0.447^{*}$ | $(0.006)$ |  |  |  |  |  |
| Unstable-Medium Adj. | $1.387^{* *}$ | $(0.013)$ | $0.439^{*}$ | $(0.007)$ |  |  |  |  |  |
| Unstable-Slow Adjust | $1.364^{* *}$ | $(0.011)$ | 0.399 | $(0.007)$ |  |  |  |  |  |

Notes: First 10 or 15 seconds omitted from each 90 or 150 second period;

* and ${ }^{* *}$ denote significant differences from Nash equilibrium at . 05 and . 01 levels for 2-tailed Wilcoxon signed-rank test, using Holm-Bonferroni p-value correction to adjust for family-wise error rate.

The right side of Table 2 reports predicted and observed dispersion as measured by the

[^5]Table 3: OLS Regressions of Summary Price Statistics

| Variable | Median Price |  | Interquartile Range |  |
| :--- | :---: | :---: | :---: | :---: |
| Unstable Market | $-0.163^{* *}$ | $(0.008)$ | -0.008 | $(0.007)$ |
| Professionals | 0.014 | $(0.015)$ | $0.045^{*}$ | $(0.020)$ |
| No Counterfactual Projection | $0.052^{* *}$ | $(0.012)$ | 0.019 | $(0.012)$ |
| No Information on Others' Prices | $0.057^{* *}$ | $(0.016)$ | $0.022^{*}$ | $(0.010)$ |
| Fast Adjustment Speed | $-0.019^{*}$ | $(0.007)$ | 0.004 | $(0.005)$ |
| Slow Adjustment Speed | $-0.018^{*}$ | $(0.007)$ | $-0.034^{* *}$ | $(0.006)$ |
| Intercept | $1.494^{* *}$ | $(0.011)$ | $0.413^{* *}$ | $(0.006)$ |
| Observations (clusters) | $216(39)$ | $212(38)$ |  |  |
| R-squared | 0.707 | 0.203 |  |  |
| Notes: Standard errors (in parentheses) adjusted for clustering on markets; * |  |  |  |  |
| and ** denote significant differences from 0 at .05 and .01 levels (2-tail tests) |  |  |  |  |

interquartile range of prices. Observed dispersion is often as predicted and sometimes smaller in the Unstable treatment; differences from equilibrium are insignificant except for some price adjustment speeds in the limited information condition. In the Stable treatment observed dispersion is consistently larger than predicted and often significantly larger. This is due mainly to the sprinkling of price choices below $\underline{p}=1.11$ and the clump at the maximum $p^{*}=1.80$. Table 8 in Appendix B shows that neither the median price nor the IQR dispersion measure change much between the first and second halves of the trading period in any treatment. Thus the time-average price distributions seem not to be trending towards a different location or degree of dispersion.

Table 3 reports OLS regressions of median prices and interquartile range, averaged for all 4 periods in each block, on a series of dummy variables for exogenous treatments. Consistent with Hypothesis I, median prices are significantly lower with the Unstable parameter vector. Relative to High information (the omitted condition), median prices are significantly higher without the counterfactual payoff projection, whether or not display of other sellers' prices is also suppressed. Median prices are slightly lower for the fast and slow adjustment speed relative to the intermediate (omitted) medium adjustment speed condition. Notably, the use of professional businesspeople as subjects has no significant impact on median prices.

The rightmost column of Table 3 shows some small differences in dispersion, with de-
creases for the slow adjustment speed and increased dispersion for professionals and when price information is suppressed. ${ }^{7}$ Dispersion is not significantly different between the Stable and Unstable market parameters, but our design deliberately set similar Nash equilibrium dispersion across the treatments ( 0.37 for Stable and 0.41 for Unstable) in order to focus on identification of differences in cycles.

In sum, we have mixed support for Hypothesis I. Nash equilibrium correctly predicts the comparative statics for median prices. Despite slightly underpredicting IQR, it is reasonably accurate in predicting the overall distribution in the $S$ treatment, at least when sellers can observe other sellers' contemporaneous prices. The NE prediction in the $U$ treatments is less accurate; although roughly correct, particularly for the level of dispersion, it significantly underpredicts the observed median price.

### 4.2 Price Cycles

Might the observed skew towards higher-than-NE prices in the $U$ treatments be due to price cycles? Figure 7 illustrates observed price dynamics in (median, IQR)-space in two trading periods. Time progresses downward, beginning at second 10 at the top of the vertical axis and ending in the horizontal plane at second 90 (or 150). The red 'pole' in the center of each figure denotes the time averaged (median, $I Q R$ ) for the period. The two panels are representative in that we see rough cycles in all speed and stability treatments, although there is considerable heterogeneity within and across treatments.

To test Hypothesis II on price cycles, we implement a version of Poincaré's technique in (median, IQR) space. Of course, in any two-dimensional space, a Poincaré section is just a line segment, i.e., a "tripwire." As detailed in Appendix B, our tripwire connects the center of the (median, IQR) observations to the horizontal axis, and each period we count the number CW of tripwire crossings in the predicted clockwise direction and the number CCW in the opposite (counterclockwise) direction. We define the Cycle Rotation Index for

[^6]

Figure 7: Examples of median-IQR dynamics in Stable-Fast treatment (Panel a) and Unstable-Slow treatment (Panel b) treatments. The straight red line plots the time averaged (median, IQR) for the period. Samples exclude first 10 seconds of each period.
each six seller market each period by

$$
\begin{equation*}
C R I=\frac{C W-C C W}{C W+C C W} \tag{7}
\end{equation*}
$$

By construction, $C R I \in[-1,1]$, with positive values indicating mostly clockwise transits. As noted earlier, the best-response cycles in median-IQR space rotate clockwise around the cycle center, so Hypothesis IIa entails positive CRI values.

Table 4 displays the average number of both kinds of transits and the average $C R I$ per period, disaggregated by treatment condition and separately for the high and limited information cases. As per Hypothesis IId, transits are less frequent when prices adjust to sellers' target prices more slowly. This is unsurprising, as the minimum time to complete a given cycle is longer in the slower adjustment treatments.

More importantly, the Cycle Rotation Index is significantly positive in every treatment condition according to the conservative Wilcoxon signed-rank test consistent with Hypothesis

Table 4: Mean Transits and Cycle Rotation Index

| High Information (Pooled) | CW <br> Transits | CCW <br> Transits | Cycle Rota- <br> tion Index |
| :--- | :---: | :---: | :---: |
| Stable-Fast Adjust | $9.47(0.69)$ | $2.68(0.28)$ | $0.59^{* *}(0.04)$ |
| Stable-Medium Adj. | $7.56(0.48)$ | $2.13(0.17)$ | $0.59^{* *}(0.04)$ |
| Stable-Slow Adjust | $4.18(0.29)$ | $1.25(0.14)$ | $0.58^{* *}(0.04)$ |
| Unstable-Fast Adjust | $8.32(0.49)$ | $1.97(0.32)$ | $0.64^{* *}(0.05)$ |
| Unstable-Medium Adj. | $6.83(0.44)$ | $1.64(0.20)$ | $0.65^{* *}(0.04)$ |
| Unstable-Slow Adjust | $3.80(0.19)$ | $1.05(0.17)$ | $0.66^{* *}(0.05)$ |
| Limited Information (Pooled) |  |  |  |
| Stable-Fast Adjust | $6.44(0.49)$ | $2.86(0.27)$ | $0.40^{* *}(0.05)$ |
| Stable-Medium Adj. | $5.67(0.29)$ | $2.81(0.26)$ | $0.39^{* *}(0.05)$ |
| Stable-Slow Adjust | $4.23(0.31)$ | $1.64(0.17)$ | $0.48^{* *}(0.04)$ |
| Unstable-Fast Adjust | $6.17(0.47)$ | $2.58(0.33)$ | $0.46^{* *}(0.06)$ |
| Unstable-Medium Adj. | $4.78(0.37)$ | $2.17(0.25)$ | $0.46^{* *}(0.05)$ |
| Unstable-Slow Adjust | $3.20(0.18)$ | $0.84(0.13)$ | $0.66^{* *}(0.05)$ |
| Notes: Standard error of the means in parentheses. Double asterisks (**) |  |  |  |
| indicate two-tailed Wilcoxon signed-rank test significantly different from |  |  |  |
| zero, at the one-percent significance level, using Holm-Bonferroni p-value |  |  |  |
| correction to adjust for family-wise error rate. |  |  |  |

IIa. Cycles are noisy, of course, and not every transit is in the predicted clockwise direction. Cycles are somewhat less frequent when sellers have limited information about their potential payoffs or other sellers' prices, but the $C R I$ averages across periods are remarkably consistent within information treatments. The $C R I$ calculated separately for the first half and second half of the trading periods are also very similar; see Table 13 in Appendix B. Thus, contrary to Hypothesis I, the price distribution remains nonstationary in all treatments. Controlling for the adjustment speed and information condition, however, the $C R I$ is always larger in the Unstable than the Stable condition. We provide a statistical test for this difference below.

Amplitude is another important aspect of price cycles. As the observed (median, IQR) moves around, its path might typically stay in a smaller or larger neighborhood of the center. We therefore also compare the cycles in terms of their amplitude, defined as the mean distance from the center at which the trajectory crosses the tripwire. Table 5 shows that average amplitude is always greater with the Unstable parameter vector than the Stable vector holding other treatments constant. The relative differences are substantial, ranging from 17

Table 5: Mean (across Periods and Markets) of Median (within Period) Cycle Amplitude

|  | Fast <br> Adjustment | Medium <br> Adjustment | Slow <br> Adjustment |
| :--- | :---: | :---: | :---: |
| Stable | $0.087(0.006)$ | $0.092(0.005)$ | $0.103(0.006)$ |
| Unstable | $0.145(0.008)$ | $0.122(0.007)$ | $0.121(0.007)$ |
| $p$-value for MW Test | $<0.001^{* *}$ | $0.003^{* *}$ | $0.050^{*}$ |
| (Stable vs. Unstable) |  |  |  |
| Limited Information (Pooled) |  |  |  |
| Stable | $0.085(0.006)$ | $0.082(0.006)$ | $0.094(0.008)$ |
| Unstable | $0.134(0.010)$ | $0.120(0.011)$ | $0.110(0.009)$ |
| $p$-value for MW Test | $<0.001^{* *}$ | $0.011^{*}$ | 0.072 |
| (Stable vs. Unstable) |  |  |  |
| Normater |  |  |  |

Notes: Standard error of the means shown in parentheses. Indicated p-values are for one-tailed tests, using Holm-Bonferroni p-value correction, and $*$ and ** highlight the five- and one-percent significance thresholds.
to $67 \%$, and the differences are highly significant except for slow adjustment (which has fewer cycles). The consistently positive cycle amplitudes in all Stable conditions enable us to reject Hypothesis IIb. The systematic, larger amplitude cycles in the Unstable condition support Hypothesis IIc. Appendix B shows that these conclusions are robust to other definitions of the cycle tripwire (Tables 11 and 12) and for an alternative definition of central tendency and dispersion (mean, standard deviation; see Tables 14 and 15).

Table 6 reports OLS regressions of these cycle measures, averaged across all 4 periods of each treatment within each session, on a dummy variable for the Unstable parameters. These estimates control for the other experimental conditions using additional dummy variables. Consistent with Hypothesis II, both cycle measures are significantly higher for Unstable. The use of business professionals lowers the $C R I$ but does not significantly affect cycle amplitude. Other than stability, the only factor significantly affecting amplitude is the availability of information about other sellers' prices. The $C R I$ is also affected by how much information is provided to sellers.

Table 6: OLS Regressions of Cycle Rotation Index and Cycle Amplitude

| Variable | Cycle Rotation Index |  | Cycle Amplitude |  |
| :--- | :---: | :---: | :---: | :---: |
| Unstable Market | $0.080^{* *}$ | $(0.020)$ | $0.034^{* *}$ | $(0.004)$ |
| Professionals | $-0.156^{* *}$ | $(0.042)$ | -0.012 | $(0.007)$ |
| No Counterfactual Projection | $-0.110^{* *}$ | $(0.037)$ | 0.003 | $(0.007)$ |
| No Information on Others' Prices | $-0.262^{* *}$ | $(0.004)$ | $-0.024^{*}$ | $(0.004)$ |
| Fast Adjustment Speed | -0.002 | $(0.003)$ | 0.009 | $(0.004)$ |
| Slow Adjustment Speed | $0.007^{*}$ | $(0.003)$ | 0.003 | $(0.004)$ |
| Intercept | $0.598^{* *}$ | $(0.026)$ | $0.094^{* *}$ | $(0.004)$ |
| Observations (clusters) | $212(38)$ |  | $212(38)$ |  |
| R-squared | 0.305 |  | 0.344 |  |
| Notes: Standard errors (in parentheses) adjusted for clustering on markets; * |  |  |  |  |
| and ** denote significant differences from 0 at .05 and .01 levels (2-tail tests) |  |  |  |  |

### 4.3 Price Changes

Recall that, although sellers could change prices at any time, for data analysis we recorded the actual price vector at 400 ms intervals, and at the same time also recorded the vector of prices targeted by each seller. The two price vectors are usually quite close together in the Medium and Fast adjustment data, but they can differ substantially when price adjustment is Slow. In this subsection, we focus on the changes in target price between successive observations.

After excluding the initial 10 or 15 adjustment seconds for the 38 markets with 6 sellers, we have $1,213,674$ adjacent 400 ms intervals to consider for possible price changes by individual sellers. Sellers changed their target price in 622,725 of these intervals, or $51 \%$. One reason for this high percentage is that our interface made price changes very easy: when sellers dragged the price slider around, as they often did, our data indicate a constantly moving target price.

Target price changes are most common in the Fast adjustment condition (59.3\%), and less frequent in the Medium (55.1\%) and Slow adjustment conditions (44.2\%). More importantly, price changes are systematically more frequent in the Unstable markets than in Stable markets. Overall, price changes occur $55.0 \%$ of the time in the Unstable periods, compared to $47.6 \%$ for the Stable periods. For 34 of 38 independent markets, target price changes are
less frequent in Stable than in Unstable periods. The relevant Wilcoxon signed-rank test establishes that the difference is significant at $p<0.001$. However, Table 16 in Appendix B shows that the frequency of price changes hardly declines from the early to late parts of the trading period in any type of market.

Considerable heterogeneity exists in the frequency of price changes across individual sellers. Only 13 of 228 sellers change prices in more than three-fourths of successive observations, and 37 of 228 sellers leave prices unchanged in more than two-thirds of successive observations. The interquartile interval for target price change frequency is [0.399, 0.645] across sellers. Average price change frequency is lower among professionals (0.488) and in markets with limited payoff and price feedback (0.476), relative to markets with student sellers and high information (0.573). This is due, in part, to professionals and sellers in limited information markets more frequently leaving prices unchanged at the maximum level.

Profits do not vary much across sellers because, as we have seen, the overall price distribution is close to the NE distribution where (within its support) profits are equal for all price choices. The lowest average earning rate among sellers is 126.3 and the highest is 164.4 , with an interquartile range [135.3, 144.2]. Despite this compressed range, we find that sellers who changed their target price frequently earned greater profits; the Spearman rank correlation of average payoff rate with price change frequency across sellers is $\rho=0.37$. This correlation is highly significant based on a cross-sectional regression that clusters standard errors at the market level ( $p<0.001$ ).

How do target prices compare to the contemporaneous best response (BR) price? More often than not, sellers move their targets in the right direction: $61 \%$ of price changes are upward when the BR exceeds the current price, and $58 \%$ are downward when it is below. There is, however, an interesting asymmetry in how often sellers keep price targets unchanged: $43 \%$ of the time when the BR exceeds the current price, but $53 \%$ of time when it is below. This relative reluctance to respond to the incentive to lower prices might be due to the fact that the local profit gradient is typically positive. That is, as can be visualized in Figure 5, a seller's counterfactual profit declines as she reduces price until her rank in the price distribution falls below a rival seller's. In other words, while simple (myopic) payoff-based learning
would facilitate directional learning (Selten and Stoecker, 1986) in the case of upward adjustments, it does not in the case of downward adjustments as subjects need to have some patience/foresight. ${ }^{8,9}$

Finally, we note that when sellers move their target towards the BR price, they tend to undershoot; see Table 17 in Appendix B. Overshooting (setting the target more than 0.05 beyond the BR , as a seller might do in anticipation that the BR will soon change) is rare, less than $5 \%$ of cases, and setting the target within 0.05 of the BR is definitely less common than setting it short of that. This behavioral inertia might suggest lack of foresight, but it might also simply be a mechanical artifact of the user interface, since most sellers drag their slider towards their true target rather than clicking on it directly.

## 5 Discussion

Our experiment investigates price dispersion in the sequential search model of Burdett and Judd (1983). That model predicts parameter-specific distributions of prices, and our aggregate data largely bear out those predictions. For "Stable" parameters, student subjects and high information treatments, the prediction is remarkably close to the empirical distribution that we observe when aggregating over sellers and over time. Predicted distributions are in the right ballpark for all treatments, and key comparative statics predictions are correct, e.g., median observed price is indeed lower with the Stable parameters.

Disaggregated data reveal much more activity than suggested by static equilibrium predictions. In all treatments, most individual sellers continually adjust prices throughout the trading period. Our data thus rule out several plausible hypotheses: that individual sellers' prices stabilize, that price dispersion fades away, and that dispersion is not systematically

[^7]related to search technology.

In our view, the most important finding is that individual sellers' price adjustments do not tend to cancel out as in some stochastic equilibrium (in which the equilibrium distribution holds at the population level, but not at the individual level); instead, there are strong and persistent cycles in the price distribution. Consistent with learning and evolutionary dynamics, we see that, starting at relatively high values, dispersion and median price both tend to fall as sellers get pulled into intense price undercutting. Dispersion continues to fall as the median price begins to stabilize at a low level. Then a few sellers abandon undercutting and jump to high prices. As more follow, the median increases sharply and dispersion also increases until the undercutting phase resumes. This process produces irregular but persistent clockwise cycles in (median, IQR) space, where IQR is a robust measure of dispersion, implying price dispersion is higher when prices are rising.

In each treatment allowing timely price adjustment, we see on average 5-10 clockwise cycles per trading period. Of course, cycles are less frequent when we enforce very gradual price adjustment (slow treatment), but the cycle rotation index still shows very significant clockwise cycling. Cycle amplitudes are significantly smaller with Stable than with Unstable parameters, but (contrary to our initial hypothesis) they remain significantly positive.

One might ask how is it possible for the overall empirical distribution to be close to the equilibrium, even while persistent cycles keep the distribution some distance away from equilibrium. To explain, consider someone constantly driving on a ring road outside of a city. An average of this driver's GPS coordinates would give the center of the ring somewhere in downtown, but the driver never was downtown, just as our subjects are never in equilibrium. Theory formally outlining this possibility is in Benaïm et al. (2009). Further, this phenomenon, persistent cycles but averages close to equilibrium, is also reported in Cason and Friedman (2003) and earlier in Brown Kruse et al. (1994).

Our results are robust in some important directions. First, to conduct an artefactual field experiment, we ran additional sessions with a different subject pool, business professionals enrolled in executive MBA programs, many of whom had field experience setting prices.

One might conjecture that such subjects would be more successful in finding equilibrium or, alternatively, in managing to collude or otherwise altering dynamic behavior. Although we detect some interesting differences - cycles are slightly more dispersed and slightly less regular - all major conclusions are confirmed when business professionals set prices.

A second robustness check was to relax our full information baseline condition, which was intended to give Nash equilibrium its best shot. Sellers in baseline always see other sellers' current prices and can easily find the best response to those prices. Sellers in many real-world situations do not have optimization tools and/or timely access to rivals' prices, so we tested the impact of going to the other extreme. One limited information treatment removed access to counterfactual payoffs (including best response), and a further treatment also removed access to other sellers' current prices. The impact turns out to be less than we expected. Under either of these limited information conditions, median prices are somewhat higher and cycles are slightly less consistent in direction and slightly lower in amplitude. However, in all settings we still observe strong cycles that show no tendency to fade away.

What might the results tell us about pricing in the wider world? Online sellers can easily and quickly adjust price offers whenever they desire, although actionable feedback typically takes hours or days to accumulate. Taking into account the different effective time scales, we see no insurmountable barrier to generalizing the theory and results presented here. Certainly field data present major challenges - e.g., in accounting for subtle heterogeneities in products and for endogenous selection of search technology - but robust results from tightly controlled settings such as ours can help isolate and cope with those additional challenges. In particular, empirical economists might consider looking for irregular cycles (perhaps using Poincaré sections) in field data. They could investigate whether the degree of dispersion, and the amplitude and frequency of cycles, varies with changes in the search technology, such as recommender systems or search engines, in a manner consistent with our laboratory findings.

Many questions also remain for future lab experiments. What happens if periods run for a very long time - would the cycle amplitude eventually fall in the Stable condition? How exploitable are the cycles - could an algorithm that tracks cycles make substantial excess
profit? Would sellers' behavior change if buyers were human subjects who could adapt their search behavior or could substitute intertemporally?

We hope that our results also provoke new theoretical developments. Appendix A raises as many questions as it answers about how economists might identify cycles in higher dimensions. Our results are broadly consistent with simple models of adaptive learning, but also highlight issues in applying existing learning models to realistically rugged payoff landscapes. For example, in our setup (and in some real world situations), a seller's payoff typically increases with small price increases, but the global best response often is to jump to a much lower price. How should economists model and identify adjustment dynamics in that sort of environment?

## Appendix A: Theory

Equilibrium. We now review some details of equilibrium calculations in BJ83 sequential search model with a continuum of sellers and a continuum of buyers. Each buyer receives $k \geq 1$ random price quotations with given probabilities $q_{k}$, with $q_{k}=0$ for $k>3$, and purchases at the lowest of these prices if at or below the endogenous reservation price $p^{*}$.

For an arbitrary cdf of seller prices $F$, each individual seller therefore faces demand

$$
\begin{equation*}
D(p)=\mu\left(q_{1}+2 q_{2}(1-F(p))+3 q_{3}(1-F(p))^{2}\right) \tag{8}
\end{equation*}
$$

on $\left[0, p^{*}\right]$ and zero elsewhere, where $\mu$ is the buyer/seller mass ratio. The mixed equilibrium distribution $F(p)$ solves $\pi(p) \equiv p D(p)=p^{*} \mu q_{1} \equiv \pi\left(p^{*}\right)$ for all $p \in\left[\underline{p}, p^{*}\right]$. That is, the profit at each price in the support interval of the equilibrium distribution is equal to the profit earned when charging the reservation price, where the seller's customers see only one price quotation. One can write $D(p)$ as $D(u)$ where $u=F(p)$ one's rank in the price distribution, giving equation (1) in the text. The equilibrium condition therefore can be rewritten as

$$
\begin{equation*}
F^{-1}(u) D(u)=p^{*} D(1) \tag{9}
\end{equation*}
$$

This implies that the inverse equilibrium price distribution has the explicit form

$$
\begin{equation*}
F^{-1}(u)=\frac{p^{*} D(1)}{D(u)} \tag{10}
\end{equation*}
$$

This leads directly to equation (2) for $F^{-1}$, and to equation (3) for $\underline{p}=F^{-1}(0)$. The reservation price $p^{*}$ equates the incremental cost of another search to the expected price reduction, i.e., it solves

$$
\begin{equation*}
c=\int_{0}^{1}\left(p^{*}-F^{-1}(u)\right) d G(u)=\int_{0}^{1} G(u) \frac{d F^{-1}(u)}{d u} d u=p^{*}\left(1-q_{1}\right), \tag{11}
\end{equation*}
$$

where $G(u)=q_{1} u+q_{2}(1-u)^{2}+q_{3}(1-u)^{3}$, the distribution of lowest prices. This gives the result that $p^{*}=c /\left(1-q_{1}\right)$.

Stability. Following HS02, we now give some results on stability under adaptive dynamics. Let the price range $[\underline{p}, \bar{p}]$ be approximated by an evenly-spaced finite grid of size $n .{ }^{10}$ That is, the strategy set is $[\underline{p}, \underline{p}+\varepsilon, \underline{p}+2 \varepsilon, \ldots, \bar{p}]=\left[p_{1}, \ldots, p_{n}\right]$ where $\varepsilon=(\bar{p}-\underline{p}) /(n-1)$. Let $x$ be a point in the $n$-simplex so that it represents the relative frequencies of strategies and let $X\left(p_{i}\right)=X_{i}=\sum_{j \leq i} x_{j}$ be the associated cumulative distribution. From this we construct a finite approximation to the BJ83 continuous profit function

$$
\begin{equation*}
\pi_{n}(x)=\left(p_{1} D\left(X_{1}\right), \ldots, p_{n} D(1)\right) \tag{12}
\end{equation*}
$$

We then consider PDA evolutionary dynamics of the form

$$
\begin{equation*}
\dot{x}=Q(x) \pi_{n}(x) \tag{13}
\end{equation*}
$$

where $Q(x)$ is a positive definite matrix and $\pi_{n}(x)$ is the vector of payoffs to the different prices. The best-known dynamics of this form are the evolutionary replicator dynamics, where the relative growth rate of frequency of strategies is proportional to relative payoffs. HS02 find that the linearization of these dynamics at a mixed equilibrium $x^{*}$ will be

$$
\begin{equation*}
Q\left(x^{*}\right) \pi_{n}^{\prime}\left(x^{*}\right) \tag{14}
\end{equation*}
$$

where $\pi_{n}^{\prime}(x)=d \pi_{n}(x) / d x$ is a $n \times n$ matrix. HS02 show that the mixed equilibrium will be asymptotically stable (unstable) under all dynamics of form (13) if $\pi_{n}^{\prime}(x)$ is negative (positive) definite. Proposition 4 in HS02 then shows that the definiteness of $\pi_{n}^{\prime}(x)$ depends on the function,

$$
\begin{equation*}
\Theta(u)=D^{\prime \prime}(u) D(u)-D^{\prime}(u)^{2}=6 q_{1} q_{3}-4 q_{2}^{2}-12 q_{2} q_{3}(1-u)-18 q_{3}(1-u)^{2} \tag{15}
\end{equation*}
$$

If $\Theta(u)>0$ (and thus $D(u)$ is $\log$ convex) for $u=0$ then $\Theta>0$ for all $u \in[0,1]$ and $\pi_{n}^{\prime}(x)$ is negative definite and the equilibrium is asymptotically stable, but if $\Theta(u)$ is negative (and

[^8]

Figure 8: Win bonus versus loss penalty.
thus $D(u)$ is $\log$ concave) for $u=1$ then $\pi_{n}^{\prime}(x)$ is positive definite. The curve in Figure 1 surrounding the Stable area is thus $\Theta(0)=0$ and $\Theta(1)=0$ bounds the Unstable area. Proposition 4 of HS02 then implies both the earlier Proposition 1 and the following.

Proposition 2. The mixed equilibrium $F_{S}$ in the game $S$ is asymptotically stable under the PDA dynamics (13), but the mixed equilibrium $F_{U}$ in the game $U$ is unstable.

To illustrate the intuition, consider $D(0.5)$, the demand from being at the midpoint of the price distribution, to be the tie payoff. Take $D(1)$ and $D(0)$ to be the loss and win payoffs from being the highest/lowest priced seller respectively. So, the win premium is $D(0)-D(0.5)$ and the loss penalty is $D(0.5)-D(1)$. Then one can see in Figure 8 that our Stable parameterization gives rise to convex demand so that the win premium is bigger than the loss penalty. However, in the unstable case, demand is linear so that they are equal.

Thus, as HS02 write, stability requires "sufficient ignorance"; that is, a large number of poorly informed buyers. The Stable game $S$ has some highly informed consumers but many poorly informed consumers. This both raises the payoff to wins (because the highly informed buyers successfully find the lowest priced seller) and reduces the size of losses (the mass of poorly informed buyers still buy from you even if you are high priced). This is what gives convexity in demand and hence stability. However, the unstable game $U$ does not have as many poorly informed consumers and so losses are quite painful.

## (Perturbed) Best Response Dynamics.

The text sometimes refers specifically to BR or Perturbed BR dynamics, which are described by the ordinary differential equation,

$$
\begin{equation*}
\dot{x}=\phi(x)-x \tag{16}
\end{equation*}
$$

where the function $\phi(\cdot)$ is a perturbed choice function such as the logit. These PBR dynamics are discussed in detail in Hopkins (1999). The fixed point $\hat{x}$ of the PBR dynamics is not in general identical to the Nash equilibrium but instead is Quantal Response (QRE) or Logit (LE) equilibrium. PBR choice functions such as the logit are usually parameterized with a precision parameter $\lambda$, which is the inverse of the noise affecting the individual's choice. Thus, the PBR dynamics approach the (exact) BR dynamics analyzed in, for example, Gaunersdorfer and Hofbauer (1995), as $\lambda$ becomes large. The results of Hopkins (1999) imply that the linearization of the PBR dynamics at $\hat{x}$ is

$$
\begin{equation*}
\lambda Q(\hat{x}) \pi_{n}^{\prime}(\hat{x})-I, \tag{17}
\end{equation*}
$$

where $\lambda$ is the precision parameter. Proposition 1 of Hopkins (1999) then implies the following: ${ }^{11}$

Proposition 3. The perturbed equilibrium (QRE) $\hat{x}$ of the game $S$ is globally asymptotically stable under the $P B R$ dynamics for all $\lambda \geq 0$.

The equivalent result for the unstable equilibrium is somewhat different. The perturbations have a stabilizing effect. The unstable equilibrium is only unstable under the perturbed dynamics if the precision is high enough, equivalent to the level of noise being sufficiently low. The value $\lambda^{*}$ is found through numerical calculation.

Proposition 4. In game $U$, the perturbed equilibrium (LE) $\hat{x}$ is unstable under the logit form of the $P B R$ dynamics for all $\lambda>\lambda^{*} \approx 0.69$.

[^9]
## Detecting Cycles.

What sort of behavior should we expect when the fixed point is unstable? In two dimensions, if the unique equilibrium point is dynamically unstable, then the classic PoincaréBendixson Theorem tells us that any sort of continuous dynamics must have a limit cycle in a generalized sense. The nature of that cycle can depend on the specific dynamics. For example, when the NE is unstable in an asymmetric Rock-Paper-Scissors game, we get convergence to a stable interior Shapley polygon under BR dynamics, but get convergence to a heteroclinic cycle (i.e., to the edges of the simplex) under replicator dynamics. The Hopf bifurcation is a textbook example of how behavior in two dimensions can transition from convergence to an equilibrium point to regular cycles as parameters change in a two dimensional system. Continuous time dynamics in systems of dimension 3 and higher can be much more complicated. Cycles can be quite intricate, and chaos (or convergence to a strange attractor) is possible.

Our dynamical system can be formalized as a system of ordinary differential equations in the infinite dimensional space $\mathcal{F}$ of probability measures (or cumulative distribution functions) whose support is contained within a fixed interval $[\underline{p}, \bar{p}]$. One can deal (as in the previous subsection) with approximations of arbitrarily large finite dimension, or (as a literal interpretation of our experiment with six sellers using pure strategies) of fixed finite dimension. In the latter case, one begins with the hypercube $[\underline{p}, \bar{p}]^{6}$, and then recognizes that the theory treats the sellers symmetrically so that only the distribution of prices matter, not the identity of the sellers. The proper domain for dynamics (or state space) thus is the $\frac{1}{6!}=1 / 320$ slice of the hypercube that satisfies the order statistics condition $\underline{p} \leq x_{1} \leq x_{2} \leq \ldots \leq x_{6} \leq \bar{p}$. That state space is still 6 dimensional, of course.

Whether the state space is infinite dimensional or 6 dimensional, it seems that very complex behavior is possible. It is also possible that behavior settles down to a nice limit cycle, plus perhaps some residual behavioral noise. To detect nice cycles, we can construct a Poincaré section of co-dimension 1, as in Figure 3 in any state space of dimension of two or larger, including infinite dimensional state spaces.

Given data consisting of observations of dispersed prices $\left(z_{1}, z_{2}, \ldots, z_{t}, \ldots, z_{T}\right)$ sampled on fine time grid, suppose that we have constructed a Poincaré section PS at point $P$ with tangent vector $N .{ }^{12}$ A positive transit $(+1)$ through PS is detected when $z_{t} \cdot N<0$ and $z_{t+1} \cdot N \geq 0$, i.e., successive dot products go from negative to positive, i.e., the trajectory passes through the section normal to the tangent vector between times $t$ and $t+1$.

A major advantage of this approach for us is that it works even if the cycle is a bit irregular or has a small but persistent stochastic component. Better known spectral techniques deal less effectively with such complications. See Wang et al. (2017) for descriptions of other possible methods such as velocity fields and angular momentum. For some purposes, our method has the disadvantage that it will not distinguish between limit cycles and some sorts of chaotic behavior (such as Lorenz attractors) that are almost periodic.

As noted in the paper, we chose for practical and metaphysical reasons to work with a two dimensional projection such as (median, IQR); Appendix B notes similar results with (mean, standard deviation.) Nevertheless, it should be acknowledged that such a low dimensional projection can be misleading. Non-cyclic behavior can project to cyclic behavior, e.g., a helix projected orthogonally to its axis looks like a cycle. Conversely, cyclic behavior can project to something that looks noncyclic, e.g., a narrow ellipse projected along its major axis looks like negligible vibration around a fixed point.

[^10]
## Appendix B: Supplementary Data Analysis

This appendix describes the steps taken to document price cycles, and also presents supplementary tables and robustness checks for the results presented in the main text.

Figures 9 and 10 reproduce the empirical CDFs for subsets of the data based on subject type (students versus professionals) and the type of feedback sellers receive. Table 2 in the main text summarizes the price data (median and interquartile range) for the entire trading period after excluding some seconds for sellers' initial adjustment. Table 7 reproduces these statistics in the columns labeled Pooled, and also subdivides the data into different seller types and the information feedback condition. Dispersion tends to be higher with professionals, but other differences do not appear to be systematic. Table 8 pools across all markets, and also subdivides the trading periods into first and second halves. Although median prices rise by one or two cents between the early and later phases of the trading period, otherwise the summary statistics are similar across the two phases.


Figure 9: Cumulative distributions of observed vs NE prices in treatments with High Information, by Subject Pool. Samples exclude first 10-15 seconds of each period.

Table 7: Mean of Summary Price Statistics (and standard errors across periods) for different information feedback and subject types

|  | Median Price |  |  | Interquartile Range |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| High Information | Pooled | Students | Professionals | Pooled | Students | Professionals |
| Nash Equilibrium | Stable) | 1.48 |  | 0.37 |  |  |
| Stable-Fast Adjust | 1.482 | 1.488 | 1.470 | 0.433 | 0.422 | 0.462 |
|  | $(0.007)$ | $(0.008)$ | $(0.015)$ | $(0.006)$ | $(0.006)$ | $(0.018)$ |
| Stable-Medium Adjust | 1.496 | 1.505 | 1.475 | 0.447 | 0.423 | 0.505 |
|  | $(0.009)$ | $(0.009)$ | $(0.019)$ | $(0.009)$ | $(0.004)$ | $(0.024)$ |
| Stable-Slow Adjust | 1.469 | 1.457 | 1.497 | 0.398 | 0.394 | 0.408 |
|  | $(0.009)$ | $(0.010)$ | $(0.017)$ | $(0.007)$ | $(0.007)$ | $(0.017)$ |
| Nash Equilibrium | (Unstable) | 1.20 |  |  | 0.41 |  |
| Unstable-Fast Adjust | 1.324 | 1.321 | 1.331 | 0.411 | 0.392 | 0.463 |
|  | $(0.008)$ | $(0.009)$ | $(0.016)$ | $(0.009)$ | $(0.007)$ | $(0.024)$ |
| Unstable-Medium Adj. | 1.339 | 1.312 | 1.404 | 0.403 | 0.397 | 0.420 |
|  | $(0.011)$ | $(0.011)$ | $(0.019)$ | $(0.008)$ | $(0.007)$ | $(0.020)$ |
| Unstable-Slow Adjust | 1.317 | 1.321 | 1.309 | 0.377 | 0.366 | 0.402 |
|  | $(0.010)$ | $(0.011)$ | $(0.019)$ | $(0.008)$ | $(0.006)$ | $(0.020)$ |
| Limited Information | Pooled | No Prices | No Pay Proj. | Pooled | No Prices | No Pay Proj. |
| Nash Equilibrium | $($ Stable $)$ | 1.48 |  |  | 0.37 |  |
| Stable-Fast Adjust | 1.528 | 1.532 | 1.525 | 0.426 | 0.444 | 0.409 |
|  | $(0.006)$ | $(0.009)$ | $(0.007)$ | $(0.007)$ | $(0.011)$ | $(0.009)$ |
| Stable-Medium Adjust | 1.550 | 1.559 | 1.533 | 0.411 | 0.416 | 0.406 |
|  | $(0.007)$ | $(0.010)$ | $(0.010)$ | $(0.006)$ | $(0.008)$ | $(0.009)$ |
| Stable-Slow Adjust | 1.548 | 1.549 | 1.546 | 0.393 | 0.381 | 0.405 |
|  | $(0.006)$ | $(0.008)$ | $(0.009)$ | $(0.006)$ | $(0.007)$ | $(0.010)$ |
| Nash Equilibrium | $($ Unstable) | 1.20 |  |  | 0.41 |  |
| Unstable-Fast Adjust | 1.356 | 1.350 | 1.361 | 0.447 | 0.450 | 0.444 |
|  | $(0.009)$ | $(0.013)$ | $(0.013)$ | $(0.006)$ | $(0.008)$ | $(0.010)$ |
| Unstable-Medium Adj. | 1.387 | 1.377 | 1.398 | 0.439 | 0.448 | 0.430 |
|  | $(0.013)$ | $(0.019)$ | $(0.018)$ | $(0.007)$ | $(0.008)$ | $(0.011)$ |
| Unstable-Slow Adjust | 1.364 | 1.374 | 1.354 | 0.399 | 0.387 | 0.411 |
|  | $(0.011)$ | $(0.016)$ | $(0.014)$ | $(0.007)$ | $(0.010)$ | $(0.011)$ |
|  |  |  |  |  |  |  |

Notes:. First 10 or 15 seconds omitted from each 90 or 150 second period.

Table 8: Mean of Summary Price Statistics (and standard errors across periods) for First and Second Halves of the Trading Periods

|  | Median Price |  |  | Interquartile Range |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| Treatment | All Period | 1st Half | 2nd Half | All Period | 1st Half | 2nd Half |
| Nash Equilibrium | (Stable) | 1.48 |  |  | 0.37 |  |
| Stable-Fast Adjust | $1.503^{*}$ | 1.497 | $1.508^{*}$ | $0.430^{* *}$ | $0.435^{* *}$ | $0.424^{* *}$ |
|  | $(0.009)$ | $(0.009)$ | $(0.009)$ | $(0.009)$ | $(0.010)$ | $(0.008)$ |
| Stable-Medium Adjust | $1.518^{* *}$ | 1.509 | $1.527^{* *}$ | $0.431^{* *}$ | $0.439^{* *}$ | $0.423^{* *}$ |
|  | $(0.010)$ | $(0.011)$ | $(0.010)$ | $(0.008)$ | $(0.009)$ | $(0.08)$ |
| Stable-Slow Adjust | 1.504 | 1.492 | $1.516^{*}$ | $0.396^{* *}$ | $0.397^{*}$ | $0.395^{* *}$ |
|  | $(0.010)$ | $(0.010)$ | $(0.011)$ | $(0.006)$ | $(0.008)$ | $(0.006)$ |
| Nash Equilibrium | (Unstable) | 1.20 |  |  | 0.41 |  |
| Unstable-Fast Adjust | $1.338^{* *}$ | $1.338^{* *}$ | $1.338^{* *}$ | 0.427 | 0.433 | 0.422 |
|  | $(0.008)$ | $(0.008)$ | $(0.010)$ | $(0.010)$ | $(0.011)$ | $(0.009)$ |
| Unstable-Medium Adj. | $1.361^{* *}$ | $1.357^{* *}$ | $1.365^{* *}$ | 0.419 | 0.420 | 0.419 |
|  | $(0.014)$ | $(0.014)$ | $(0.014)$ | $(0.008)$ | $(0.009)$ | $(0.009)$ |
| Unstable-Slow Adjust | $1.338^{* *}$ | $1.330^{* *}$ | $1.347^{* *}$ | 0.387 | 0.382 | 0.392 |
|  | $(0.011)$ | $(0.010)$ | $(0.012)$ | $(0.009)$ | $(0.010)$ | $(0.008)$ |

Notes:. First 10 or 15 seconds omitted from each 90 or 150 second period; * and ${ }^{* *}$ denote significant differences from Nash equilibrium at .05 and .01 levels for 2-tailed Wilcoxon signed-rank test, using Holm-Bonferroni p-value correction to adjust for family-wise error rate.


Figure 10: Cumulative distributions of observed vs NE prices in treatments with Limited Information. Samples exclude first 10 seconds of each period.

In Section 4.2, we implement the Poincaré procedure for each group of six sellers and each market period as follows.

1. Sample. Observe the vector of subjects' posted prices at 400 ms intervals, beginning at second 10 (or 15) and ending at second 90 (or 150).
2. Two-dimensional summary. For each observation $t$, compute the two-dimensional summary statistic $z_{t}$, the median and interquartile range (IQR) of the posted price vector.
3. Center. Compute the mean of $z_{t}$. This time-average defines the Center $C=\left(C_{m e d}, C_{i q r}\right)$ about which we may observe cycling.
4. Poincaré Sections. Construct the Vertical Tripwire as the line segment connecting $C$ to the point $V=\left(C_{m e d}, 0\right)$ on the horizontal (median price) axis. To check robustness, also construct the Horizontal Tripwire as the line segment connecting $C$ to the point
$H=\left(0, C_{i q r}\right)$ on the vertical (IQR) axis. See Appendix A for a general discussion of Poincaré sections.
5. Counting Transits. A transit occurs when the line segment connecting successive $z_{t}$ 's intersects a tripwire. If the crossing is clockwise (e.g., for the vertical tripwire, if $z_{t} \geq 0>z_{t+1}$ ), then increment CW by 1 ; otherwise increment the counter-clockwise count CCW by 1 .
6. Cycle Rotation Index. This index is defined, following Cason et al. (2014), for each period and market as the difference in the number of $C W$ transits and the number of $C C W$ transits, as a fraction of total transits in the period.
7. Cycle Amplitude. At every point where the price median and interquartile range trajectory crosses a tripwire in the predicted clockwise direction, record the Euclidean distance from the cycle center. For each period and market, define the cycle amplitude as the median recorded distance.

Tables 4 and 5 in the main text report the average number of cycle transits, the cycle rotation index, and the average cycle amplitudes based on the vertical Poincaré section tripwire, defined as the line segment connecting the empirical cycle center to the point $V=\left(C_{m e d}, 0\right)$ on the horizontal (median price) axis. Tables 9 and 10 report these statistics disaggregated by seller type and information feedback. The lower cycle rotation index when sellers receive no information about other sellers' prices indicates that cycles are less regular in this information condition.

To check robustness of the conclusions to an alternative tripwire definition, Tables 11 and 12 report cycle statistics pooled over all markets, with those based on the vertical tripwire shown on the left, and on the right their counterparts based on the horizontal tripwire, the line segment connecting $C$ to the point $H=\left(0, C_{i q r}\right)$ on the vertical (IQR) axis. Cycle amplitude for the horizontal tripwire is again greater in the Unstable than the Stable case, although the differences are smaller and they are only marginally significant for two of the three adjustment speeds.

The main difference between the tripwires is that Horizontal sees consistently more CW transits as well as more CCW transits. This reduces the CRI somewhat, but it always remains very highly significant. This difference may be due to slower movements and/or more noise in the cycle phase near the horizontal tripwire. That is, there seems to be relatively less noise in the cycle phase when median price decreases at low dispersion (near the vertical tripwire) than at the phase when dispersion increases at middling median price (near the horizontal tripwire).

As noted in the text, the cycle frequency and $C R I$ are similar in the early and late halves of each trading period, providing no evidence of convergence even for the Stable environment. This is documented in Table 13 for the Vertical Tripwire. Comparing each treatment condition across rows does not suggest any systematic change between the early and the late halves, for either the Stable or Unstable markets.

The cycles are constructed using robust measures of central tendency (median) and dispersion (interquartile range). As a further robustness check, Tables 14 and 15 reproduce the $C R I$ and cycle amplitude calculations based on an alternative two dimensions of central tendency and dispersion: price mean and standard deviation. Comparison with Tables 11 and 12 indicates that all of our conclusions regarding cycles are robust to this alternative. Although the number of transits is slightly lower (Table 14), the CRI is very similar for both versions. Cycle amplitudes are lower in the (mean, standard deviation) space (Table 15), but the key conclusion continues to hold: cycles are systematically larger in the Unstable than the Stable case.

Table 9: Mean Transits and Cycle Rotation Indexes, for different information feedback and subject types

|  | Students |  |  | Professionals |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
|  | CW | CCW | Cycle Rota- | CW | CCW | Cycle Rota- |
| High Information | Transits | Transits | tion Index | Transits | Transits | tion Index |
| Stable-Fast Adjust | 9.66 | 2.48 | 0.61 | 8.95 | 3.25 | 0.51 |
|  | $(0.71)$ | $(0.3)$ | $(0.04)$ | $(1.86)$ | $(0.69)$ | $(0.08)$ |
| Stable-Medium Adj. | 8.20 | 1.96 | 0.65 | 6.08 | 2.50 | 0.44 |
|  | $(0.47)$ | $(0.18)$ | $(0.03)$ | $(0.95)$ | $(0.33)$ | $(0.08)$ |
| Stable-Slow Adjust | 4.05 | 1.14 | 0.60 | 4.55 | 1.55 | 0.53 |
|  | $(0.31)$ | $(0.16)$ | $(0.04)$ | $(0.73)$ | $(0.27)$ | $(0.10)$ |
| Unstable-Fast Adjust | 8.89 | 1.66 | 0.70 | 6.70 | 2.85 | 0.47 |
|  | $(0.56)$ | $(0.21)$ | $(0.03)$ | $(0.60)$ | $(1.01)$ | $(0.14)$ |
| Unstable-Medium Adj. | 7.11 | 1.55 | 0.68 | 6.05 | 1.90 | 0.56 |
|  | $(0.52)$ | $(0.24)$ | $(0.04)$ | $(0.81)$ | $(0.43)$ | $(0.08)$ |
| Unstable-Slow Adjust | 3.82 | 0.84 | 0.72 | 3.75 | 1.54 | 0.53 |
|  | $(0.26)$ | $(0.2)$ | $(0.05)$ | $(0.26)$ | $(0.23)$ | $(0.09)$ |
|  |  | $\underline{\text { No Prices }}$ |  | No Payoff Projection |  |  |
|  | CW | CCW | Cycle Rota- | CW | CCW | Cycle Rota- |
|  | Transits | Transits | tion Index | Transits | Transits | tion Index |
| Limited Information | 5.81 | 3.47 | 0.27 | 7.06 | 2.25 | 0.53 |
| Stable-Fast Adjust | $(0.65)$ | $(0.39)$ | $(0.07)$ | $(0.70)$ | $(0.25)$ | $(0.05)$ |
|  | 3.16 | 3.06 | 0.28 | 6.19 | 2.56 | 0.50 |
| Stable-Medium Adj. | $(0.37)$ | $(0.30)$ | $(0.06)$ | $(0.38)$ | $(0.42)$ | $(0.07)$ |
|  | 3.59 | 1.50 | 0.47 | 4.88 | 1.78 | 0.50 |
| Stable-Slow Adjust | $(0.18)$ | $(0.19)$ | $(0.06)$ | $(0.52)$ | $(0.29)$ | $(0.06)$ |
| Unstable-Fast Adjust | 6.50 | 3.41 | 0.32 | 5.84 | 1.75 | 0.60 |
|  | $(0.68)$ | $(0.43)$ | $(0.07)$ | $(0.68)$ | $(0.31)$ | $(0.05)$ |
| Unstable-Medium Adj. | 4.50 | 2.59 | 0.34 | 5.06 | 1.75 | 0.58 |
|  | $(0.51)$ | $(0.38)$ | $(0.05)$ | $(0.55)$ | $(0.28)$ | $(0.06)$ |
| Unstable-Slow Adjust | 3.16 | 0.72 | 0.71 | 3.25 | 0.97 | 0.60 |
|  | $(0.29)$ | $(0.14)$ | $(0.05)$ | $(0.25)$ | $(0.21)$ | $(0.07)$ |

Notes: Standard error of the means shown in parentheses.

Table 10: Mean (across Periods and Markets) of Median (within Period) Cycle Amplitude, for different information feedback and subject types

|  |  | Students |  |  | Professionals |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| High Information | Fast Adj. | Med. Adj. | Slow Adj. | Fast Adj. | Med. Adj. | Slow Adj. |  |
| Stable | 0.087 | 0.092 | 0.107 | 0.086 | 0.098 | 0.088 |  |
|  | $(0.007)$ | $(0.005)$ | $(0.007)$ | $(0.008)$ | $(0.006)$ | $(0.006)$ |  |
| Unstable | 0.147 | 0.130 | 0.125 | 0.142 | 0.094 | 0.109 |  |
|  | $(0.007)$ | $(0.006)$ | $(0.005)$ | $(0.021)$ | $(0.011)$ | $(0.018)$ |  |
|  |  | No Prices | No Payoff Projection |  |  |  |  |
| Limited Information | Fast Adj. | Med. Adj. | Slow Adj. | Fast Adj. | Med. Adj. | Slow Adj. |  |
| Stable | 0.074 | 0.077 | 0.74 | 0.097 | 0.086 | 0.114 |  |
|  | $(0.003)$ | $(0.006)$ | $(0.005)$ | $(0.009)$ | $(0.011)$ | $(0.012)$ |  |
| Unstable | 0.118 | 0.107 | 0.095 | 0.150 | 0.133 | 0.125 |  |
|  | $(0.011)$ | $(0.014)$ | $(0.014)$ | $(0.015)$ | $(0.017)$ | $(0.008)$ |  |

Notes: Standard error of the means shown in parentheses.

Table 11: Mean Transits and Cycle Rotation Indexes, Pooled Markets

|  | Vertical Tripwire |  |  | Horizontal Tripwire |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
|  | CW | CCW | Cycle Rota- | CW | CCW | Cycle Rota- |
| Treatment | Transits | Transits | tion Index | Transits | Transits | tion Index |
| Stable-Fast Adjust | 8.06 | 2.76 | $0.50^{* *}$ | 9.49 | 4.16 | $0.40^{* *}$ |
|  | $(0.50)$ | $(0.20)$ | $(0.03)$ | $(0.49)$ | $(0.31)$ | $(0.04)$ |
| Stable-Medium Adj. | 6.72 | 2.43 | $0.50^{* *}$ | 7.37 | 3.10 | $0.44^{* *}$ |
|  | $(0.33)$ | $(0.16)$ | $(0.04)$ | $(0.38)$ | $(0.24)$ | $(0.04)$ |
| Stable-Slow Adjust | 4.21 | 1.43 | $0.54^{* *}$ | 4.53 | 1.57 | $0.54^{* *}$ |
|  | $(0.21)$ | $(0.11)$ | $(0.03)$ | $(0.24)$ | $(0.13)$ | $(0.03)$ |
| Unstable-Fast Adjust | 7.34 | 2.25 | $0.56^{* *}$ | 10.26 | 5.04 | $0.36^{* *}$ |
|  | $(0.38)$ | $(0.23)$ | $(0.04)$ | $(0.41)$ | $(0.32)$ | $(0.03)$ |
| Unstable-Medium Adj. | 5.89 | 1.89 | $0.56^{* *}$ | 7.14 | 3.04 | $0.45^{* *}$ |
|  | $(0.34)$ | $(0.16)$ | $(0.03)$ | $(0.32)$ | $(0.21)$ | $(0.03)$ |
| Unstable-Slow Adjust | 3.53 | 0.96 | $0.66^{* *}$ | 4.28 | 1.42 | $0.57^{* *}$ |
|  | $(0.14)$ | $(0.11)$ | $(0.03)$ | $(0.19)$ | $(0.12)$ | $(0.03)$ |

Notes: Standard error of the means in parentheses. Double asterisks ( ${ }^{* *}$ ) indicate two-tailed Wilcoxon signed-rank test significantly different from zero, at the one-percent significance level, using Holm-Bonferroni p-value correction to adjust for family-wise error rate.

Table 12: Mean (across Periods and Markets) of Median (within Period) Cycle Amplitude, Pooled Markets

|  | Crossing Vertical Tripwire |  |  | Crossing Horizontal Tripwire |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Treatment | Fast Adj. | Med. Adj. | Slow Adj. | Fast Adj. | Med. Adj. | Slow Adj. |  |
| Stable | 0.086 | 0.088 | 0.098 | 0.110 | 0.106 | 0.097 |  |
|  | $(0.004)$ | $(0.004)$ | $(0.005)$ | $(0.004)$ | $(0.004)$ | $(0.003)$ |  |
| Unstable | 0.140 | 0.120 | 0.116 | 0.126 | 0.117 | 0.109 |  |
|  | $(0.006)$ | $(0.006)$ | $(0.005)$ | $(0.006)$ | $(0.005)$ | $(0.005)$ |  |
| $p$-value for |  |  |  |  |  |  |  |
| MW Test | $<0.001^{* *}$ | $<0.001^{* *}$ | $0.018^{*}$ | $0.008^{* *}$ | 0.098 | 0.067 |  |

Notes: Standard error of the means shown in parentheses. Indicated p-values are for onetailed tests, using Holm-Bonferroni p-value correction, and * and ${ }^{* *}$ highlight the five- and one-percent significance thresholds.

Table 13: Mean Transits and Cycle Rotation Indexes, by Period Halves

|  | 1st Half |  |  |  | 2nd Half |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| Treatment | CW <br> Transits | CCW <br> Transits | Cycle Rota- <br> tion Index | CW <br> Transits | CCW <br> Transits | Cycle Rota- <br> tion Index |
| Stable-Fast Adjust | 3.95 | 1.26 | $0.58^{* *}$ | 4.13 | 1.50 | $0.49^{* *}$ |
|  | $(0.24)$ | $(0.15)$ | $(0.04)$ | $(0.26)$ | $(0.13)$ | $(0.05)$ |
| Stable-Medium Adj. | 3.18 | 1.13 | $0.54^{* *}$ | 3.54 | 1.30 | $0.50^{* *}$ |
|  | $(0.18)$ | $(0.08)$ | $(0.03)$ | $(0.20)$ | $(0.11)$ | $(0.04)$ |
| Stable-Slow Adjust | 2.14 | 0.81 | $0.54^{* *}$ | 2.06 | 0.62 | $0.63^{* *}$ |
|  | $(0.12)$ | $(0.08)$ | $(0.04)$ | $(0.13)$ | $(0.07)$ | $(0.04)$ |
| Unstable-Fast Adjust | 3.38 | 0.96 | $0.62^{* *}$ | 3.96 | 1.29 | $0.57^{* *}$ |
|  | $(0.21)$ | $(0.11)$ | $(0.05)$ | $(0.22)$ | $(0.15)$ | $(0.05)$ |
| Unstable-Medium Adj. | 2.88 | 0.96 | $0.60^{* *}$ | 3.01 | 0.93 | $0.59^{* *}$ |
|  | $(0.20)$ | $(0.11)$ | $(0.04)$ | $(0.19)$ | $(0.10)$ | $(0.04)$ |
| Unstable-Slow Adjust | 1.82 | 0.53 | $0.66^{* *}$ | 1.71 | 0.42 | $0.72^{* *}$ |
|  | $(0.18)$ | $(0.08)$ | $(0.05)$ | $(0.19)$ | $(0.14)$ | $(0.06)$ |

Notes: Standard error of the means in parentheses. Double asterisks (**) indicate two-tailed Wilcoxon signed-rank test significantly different from zero, at the one-percent significance level, using Holm-Bonferroni p-value correction to adjust for family-wise error rate.

Table 14: Mean Transits and Cycle Rotation Indexes based on Price Mean and Standard Deviation

|  | Vertical Tripwire |  |  |  | Horizontal Tripwire |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | CW | CCW | Cycle Rota- | CW | CCW | Cycle Rota- |  |
| Treatment | Transits | Transits | tion Index | Transits | Transits | tion Index |  |
| Stable-Fast Adjust | 7.12 | 2.69 | $0.47^{* *}$ | 8.13 | 3.69 | $0.40^{* *}$ |  |
|  | $(0.49)$ | $(0.22)$ | $(0.04)$ | $(0.45)$ | $(0.25)$ | $(0.03)$ |  |
| Stable-Medium Adj. | 5.35 | 1.93 | $0.50^{* *}$ | 6.34 | 2.95 | $0.41^{* *}$ |  |
|  | $(0.29)$ | $(0.15)$ | $(0.03)$ | $(0.28)$ | $(0.21)$ | $(0.03)$ |  |
| Stable-Slow Adjust | 3.18 | 1.09 | $0.54^{* *}$ | 3.94 | 1.63 | $0.47^{* *}$ |  |
|  | $(0.17)$ | $(0.09)$ | $(0.03)$ | $(0.19)$ | $(0.11)$ | $(0.03)$ |  |
| Unstable-Fast Adjust | 6.21 | 1.96 | $0.55^{* *}$ | 8.93 | 4.32 | $0.37^{* *}$ |  |
|  | $(0.38)$ | $(0.20)$ | $(0.04)$ | $(0.33)$ | $(0.30)$ | $(0.04)$ |  |
| Unstable-Medium Adj. | 4.60 | 1.25 | $0.63^{* *}$ | 5.93 | 2.51 | $0.46^{* *}$ |  |
|  | $(0.27)$ | $(0.13)$ | $(0.04)$ | $(0.226)$ | $(0.20)$ | $(0.03)$ |  |
| Unstable-Slow Adjust | 2.58 | 0.81 | $0.61^{* *}$ | 3.44 | 1.40 | $0.49^{* *}$ |  |
|  | $(0.12)$ | $(0.09)$ | $(0.04)$ | $(0.16)$ | $(0.09)$ | $(0.03)$ |  |

Notes: Standard error of the means in parentheses. Double asterisks $\left(^{* *}\right)$ indicate two-tailed Wilcoxon signed-rank test significantly different from zero, at the one-percent significance level, using Holm-Bonferroni p-value correction to adjust for family-wise error rate.

Table 15: Mean (across Periods and Markets) of Median (within Period) Cycle Amplitude, based on Price Mean and Standard Deviation

|  | Crossing Vertical Tripwire |  |  | Crossing Horizontal Tripwire |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Treatment | Fast Adj. | Med. Adj. | Slow Adj. | Fast Adj. | Med. Adj. | Slow Adj. |  |
| Stable | 0.034 | 0.034 | 0.032 | 0.072 | 0.064 | 0.057 |  |
|  | $(0.002)$ | $(0.003)$ | $(0.002)$ | $(0.002)$ | $(0.003)$ | $(0.003)$ |  |
| Unstable | 0.068 | 0.057 | 0.053 | 0.089 | 0.088 | 0.083 |  |
|  | $(0.004)$ | $(0.004)$ | $(0.004)$ | $(0.004)$ | $(0.005)$ | $(0.005)$ |  |
| $p$-value for |  |  |  |  |  |  |  |
| MW Test | $<0.001^{* *}$ | $<0.001^{* *}$ | $<0.001^{* *}$ | $<0.001^{* *}$ | $0.001^{* *}$ | $<0.001^{* *}$ |  |

Notes: Standard error of the means shown in parentheses. Indicated p-values are for onetailed tests, using Holm-Bonferroni p-value correction, and * and ${ }^{* *}$ highlight the five- and one-percent significance thresholds.

Table 16: Price Change Frequency (Percent)

|  | Throughout <br> Trading Period | First Half of <br> Trading Period | Second Half of <br> Trading Period |
| :--- | :---: | :---: | :---: |
| All Treatments | 51.3 | 51.6 | 51.0 |
| Fast Adjustment Speed | 59.3 | 59.1 | 59.4 |
| Medium Adjustment Speed | 55.1 | 55.4 | 54.7 |
| Slow Adjustment Speed | 44.2 | 44.8 | 43.6 |
| Unstable Markets | 55.0 | 55.0 | 54.9 |
| Stable Markets | 47.6 | 48.2 | 47.0 |

Note: All percentages exclude the first 10 or 15 seconds of the trading period.

Table 17: Price Target Adjustments (Percentages of Cases)

|  |  | Towards BR |  |  |
| :--- | :---: | :---: | :---: | :---: |
|  | Away from BR | Small Towards BR | Near BR | Beyond BR |
| Fast Adjustment Speed |  |  |  |  |
| BR above Current Price | $37.5 \%$ | $44.8 \%$ | $13.4 \%$ | $4.4 \%$ |
| BR below Current Price | $39.8 \%$ | $36.3 \%$ | $19.5 \%$ | $4.4 \%$ |
| Medium Adjustment Speed |  |  |  |  |
| BR above Current Price | $37.4 \%$ | $45.5 \%$ | $13.2 \%$ | $3.9 \%$ |
| BR below Current Price | $40.6 \%$ | $36.8 \%$ | $18.4 \%$ | $4.3 \%$ |

Notes: Small Towards BR denotes movements in the direction of the best response but at least five cents short of it; Near BR denotes target within 5 cents of the best response; Beyond BR denotes movements at least 5 cents beyond the best response. This table only considers the Fast and Medium adjustment speeds, since the actual price often lags considerably behind the target price with Slow adjustment.

Table 18: Treatment Condition Ordering for Each Session (39 total markets)

| Block Number (Each Block Consists of 4 Trading Periods) |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Session Label | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 |
| 101817 | U-CS | S-CI | U-CM | S-D | U-CI | S-CS | U-D | S-CM |
| 110117A | S-CS | U-CM | S-D | S-CM | U-CS | S-CI | U-D | U-CI |
| 110117B | U-CI | S-CS | S-CI | U-D | U-CM | S-D | U-CS | S-CM |
| 021918A | U-D | U-CM | S-CS | S-CI | U-CI | U-CS | S-D | S-CM |
| 021918B | S-CI | S-CS | U-D | U-CM | S-CM | U-CI | U-CS | S-D |
| 022018A | S-CS | U-CM | U-CS | U-D | S-D | S-CM | U-CI | S-CI |
| 022018B | U-D | S-CS | S-D | S-CI | S-CM | U-CI | U-CM | U-CS |
| 060119A | U-CI | S-CM | U-CS | S-CI |  |  |  |  |
| 060119B* | S-CI | U-CM | S-CS | U-CI |  |  |  |  |
| 073019A | U-CI | S-CS | U-CM | S-CI |  |  |  |  |
| 073019B | S-CM | U-CS | U-CM | S-CS |  |  |  |  |
| 082419A* | U-CM | S-CI | S-CM | U-CS |  |  |  |  |
| 082419B* | S-CM | U-CI | U-CS | S-CS |  |  |  |  |
| 011720 | U-CS | S-D | U-CM | S-CI | U-CI | S-CS | U-D | S-CM |
| 012320A | U-D | U-CM | S-CS | S-CI | U-CI | U-CS | S-D | S-CM |
| 012320B | S-D | S-CM | U-CS | U-CI | S-CI | S-CS | U-D | U-CM |
| 012320C | U-D | U-CM | S-CS | S-CI | U-CI | U-CS | S-D | S-CM |
| 012320D | S-D | S-CM | U-CS | U-CI | S-CI | S-CS | U-D | U-CM |
| 020620 | S-CI | S-CS | U-D | U-CM | S-CM | U-CI | U-CS | S-D |
| 020720A | S-D | U-CM | S-CI | U-CS | U-CI | S-CS | U-D | S-CM |
| 020720B | U-D | S-CM | U-CI | S-CS | S-CI | U-CS | S-D | U-CM |
| Notes: Each session configuration included two markets of six sellers each, except for those indicated with * that had only 1 market. Markets with business professionals only had 4 blocks ( 16 periods) U denotes Unstable, S denotes Stable, and CS, CM, CI and D denote Continuous Slow Price Adjustment, Continuous Medium Price Adjustment, Continuous Fast Price Adjustment, and Discrete Pricing Rounds, respectively. |  |  |  |  |  |  |  |  |

## Appendix C: Instructions to Subjects

Welcome! This is an economics experiment. If you pay close attention to these instructions, you can earn a significant amount of money, which will be paid to you in cash at the end of the last round.

Please remain silent, turn off your cell phone, and do not look at other participants' screens. If you have any questions, or need assistance of any kind, please raise your hand and we will come to you. If you disrupt the experiment by talking, laughing, etc., you may be asked to leave and may not be paid. We expect and appreciate your cooperation today.

## The Basic Idea

The experiment will be divided into a number of rounds and throughout all rounds you will be anonymously matched with the same five other "counterparts" via the computer. These counterparts are other participants in today's experiment.

You and your counterparts are sellers of a fictitious good, and throughout each round you will select prices to offer to the market. The buyers in today's experiment are simulated by computerized "robots." Your price choices and those of the other sellers determine your rewards for the round. Your production costs are 0 , so every sale you make directly adds to your reward, i.e., to the amount of money you will receive at the end of the experiment.

Some trading rounds will be further divided into a number of subperiods. In each subperiod you and the other sellers will privately select prices and at the end of the subperiod the combination of these prices will determine your rewards for the subperiod.

## How Buyers Purchase

Buyers want to buy at the lowest price they can, and they will never pay more than $\$ 1.80$ per unit. Buyers do not see all six sellers' prices, however. Sometimes they see one price, and as long as that price is no higher than $\$ 1.80$ they will purchase at that price. Sometimes buyers see two or three prices, and in those cases they will purchase the lowest-price item they see.

We will not tell you exactly how often buyers see one, two or three prices, or their cost to search for additional prices, but here are a few facts:

1. Your reward at any time depends entirely on your current price and the other sellers' current prices, and nothing else.
2. The function that determines how your sales quantity (and therefore your rewards) depends on your prices and others' prices will not change over the course of a round. That is, if you and the other sellers choose the same prices at time A as at time B, you will all have the same profits at time A as at time B .
3. The function that determines your sales quantity from the prices chosen can change in different rounds (due to changes in buyers' search costs or how many prices they see).
4. Your rewards are symmetric with the other sellers. That is, the function that determines how prices determine sales quantity is the same for all participants.

## The screen display

Figure 1 shows the computer display you will use to choose prices and interact with the other sellers. You will receive rewards throughout the round that depend on your price choice and other sellers' price choices, and your rewards will be displayed on this screen.

You select your price by moving a slider (from left to right) or clicking at the bottom of the screen below the left graph. This determines a "price target" that you want to offer. Your actual price being offered to buyers may quickly or slowly adjust to this target, and this adjustment speed may differ across trading rounds. The horizontal axis covers prices from $\$ 0.00$ (the far left) to $\$ 1.80$ (the far right).

If the round is divided into subperiods, you can freely adjust your tentative price offer by clicking on the screen or dragging the slider. Your actual price for the subperiod is the location of your slider at the end of the subperiod. Subperiod endings are designated with vertical lines on the rewards time graph at the right of the screen (see Figure 2).

During the round (or when the subperiod is over) you will see a green circle displaying your
rewards payoff rate at that time. The higher the circle, the faster your reward accumulates. You will also be shown circles of different colors at the bottom of the screen showing the other sellers' prices.

In some rounds your display may also show the rewards you could receive at every possible price, based on the current (or previous subperiod) prices of all other sellers. The height of this light gray line indicates the possible rewards amount at other prices you could have chosen.

It is important to keep in mind for rounds divided into subperiods that the other sellers' prices and your reward circle height always refers to outcomes from the previous subperiod. You will not learn your rewards or others' prices from the current subperiod until after the subperiod is over.

The small graph on the top right of the screen shows the history of your price choices (green line) and the other sellers' price choices (other colored lines) throughout the current trading round. Higher lines correspond to higher prices.

## Rewards

To the right of your price choice graph you will have a display showing your accumulating rewards for the current round. Your rewards are represented by the solid green area - the larger the area, the greater your accumulated rewards. The height of the green area is based on the prices chosen by you and others at that moment (or in that subperiod). So the higher the green line, the faster your rewards are accumulating.

Your rewards will be given in points. Your points will accumulate over the course of the experiment. The upper left of your screen will always display your "Round Rewards" during the round so far and "Accumulated Rewards" earned over all previous rounds.

You will be paid cash for points earned at a conversion rate written on the white board at the front of the room.


## Frequently asked questions

Q1. Is this some kind of psychological experiment with an agenda you haven't told us?
Answer. No. It is an economics experiment. If we do anything deceptive or don't pay you cash as described then you can complain to the campus Human Subjects Committee and we will be in serious trouble. These instructions are meant to clarify how the market works and show you how you earn money; our interest is simply in seeing how people make decisions.

Q2. What changes from one round to the next?

Answer. The likelihood that the robot buyers see one, two or three prices might change, and their cost to search for more prices might change, both of which will change the function that determines your sales quantity from the prices chosen. The adjustment speed might also be faster or slower than in the previous round. Also, some rounds may be broken down into subperiods. But often nothing changes between rounds, and you are matched with exactly the same people in exactly the same way as in the last round.


Figure 2 (Subperiods)

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Round Reward: 94.39
Accumulated Rewards: 650.42
Game Type: A


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    ${ }^{\dagger}$ respectively: Economics Department, Purdue University; Economics Departments, University of Essex and University of California Santa Cruz; and School of Economics, University of Edinburgh.

[^1]:    ${ }^{1}$ Zhijian Wang (private communication) reminds us that in the last 25 years, since the Fluctuation Theorem became widely recognized, physical scientists also take a more statistical approach to describing cycles.

[^2]:    ${ }^{2}$ Given the results of Baye et al. (1992) for a related model, we suspect that the 6 -player game also has numerous asymmetric NE whose composite distributions approximate $F_{U}(p)$ (or $F_{S}(p)$ ). Since our focus is the predictive power of the original BJ83 model, we will not pursue that conjecture.

[^3]:    ${ }^{3}$ These professionals were enrolled in Executive or Weekend MBA programs at Purdue University. One market with professionals had only five sellers.

[^4]:    ${ }^{4}$ All 16 of these limited information markets began with 4 periods of the high information display so that sellers could learn how their payoffs depend on all sellers' prices. We do not analyze these "training" periods.

[^5]:    ${ }^{5}$ Table 7 and Figures 9 and 10 in Appendix B summarize data disaggregated by subject pool and information condition.
    ${ }^{6}$ These and other unconditional tests in this study employ Holm-Bonferroni p-value correction, as in Andreoni et al. (2019), to adjust for the family-wise error rate (e.g., see List et al., 2019).

[^6]:    ${ }^{7}$ To avoid redefining dispersion, we exclude from the interquartile range regressions the one market with 5 rather than 6 sellers; thus these regressions use 212 rather than 216 observations.

[^7]:    ${ }^{8}$ The asymmetry in the rate of price changes when the BR is above rather than below the current price is larger in the Slow and Medium adjustment conditions (11\%) than with Fast adjustment (8\%), which is consistent with this patience/foresight conjecture. With Fast price adjustment, patience is not required since movement to the (recent) BR is nearly instantaneous.
    ${ }^{9}$ See Nax and Pradelski (2015), Nax et al. (2016) for recent theory and experimental evidence of myopic directional learning.

[^8]:    ${ }^{10}$ It is possible to extend these results to allow for prices below $p$ and or above $\bar{p}$ at the cost of additional notation. However, as these prices are dominated strategies, the population shares for these strategies would quickly approach zero.

[^9]:    ${ }^{11}$ See also Lahkar (2011).

[^10]:    ${ }^{12}$ In the infinite dimensional case, each $z_{s}$ is a probability measure with $\operatorname{cdf} F$, so the tangent vector $N$ is defined by some test function $\phi$ with integral 0 on the price interval $\left[0, p^{*}\right]$. The inner product thus is the Stieltjes integral $z_{t} \cdot N=\int_{0}^{p^{*}} \phi(x) F(d x, t)$.

